# DSGE Modeling and Statistical Inference

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- Estimated dynamic stochastic general equilibrium (DSGE) models are now widely used for
  - empirical research in macroeconomics;
  - quantitative policy analysis and prediction at central banks.
- This mini course will focus on
  - econometric methods to conduct quantitative analysis with DSGE models;
- We will start with a prototypical New Keynesian DSGE model...

- A small-scale DSGE model: specification, steady states, log-linearization, first-order approximation to equilibrium dynamics, state-space representation.
- 2 Statistical inference: frequentist versus Bayesian; use the Kalman filter to evaluate likelihood function.
- Frequentist inference: maximum likelihood, simulated minimum distance approaches, GMM
- Bayesian inference: priors, posteriors, Metropolis-Hastings algorithm, post-processing draws.
- **5** Applications to monetary and fiscal policy analysis.

# Prototypical Applications to Monetary and Fiscal Policy Analysis

- What is the optimal target inflation rate?
- 2 Was high inflation and output volatility in the 1970s due to loose monetary policy?
- 3 Effects of the zero lower bound on nominal interest rates on monetary policy.
- How large are government spending multipliers?

The model consists of

- households;
- final goods producing firms;
- intermediate goods producing firms;
- central bank and fiscal authority;
- exogenous shock processes

• Households maximize

$$\mathbb{E}_{\tau}\left[\sum_{t=\tau}^{\infty}\beta^{(t-\tau)}\left\{\ln C_{t}-\frac{\phi_{t}}{1+\nu}L_{t}^{1+\nu}\right\}\right]$$

• subject to the constraints:

$$P_tC_t + B_{t+1} \leq P_tW_tL_t + \Pi_t + R_{t-1}B_t - T_t + \Omega_t.$$

- In a nutshell:
  - household cares about the future: intertemporal optimization
  - household likes consumption
  - household does not like to work...
  - there is a budget constraint: can't spend more than you earn and borrow; have to pay taxes;

• Households maximize

$$\mathbb{E}_{\tau}\left[\sum_{t=\tau}^{\infty}\beta^{(t-\tau)}\left\{\ln C_{t}-\frac{\phi_{t}}{1+\nu}L_{t}^{1+\nu}\right\}\right]$$

• subject to the constraints:

$$P_tC_t + B_{t+1} \leq P_tW_tL_t + \Pi_t + R_{t-1}B_t - T_t + \Omega_t.$$

- Introduce Lagrange multiplier  $\mu_t$  for budget constraint.
- Lagrangian

$$\mathcal{L} = \mathbb{E}_{\tau} \left[ \sum_{t=\tau}^{\infty} \beta^{(t-\tau)} \left\{ \ln C_t - \frac{\phi_t}{1+\nu} L_t^{1+\nu} - \mu_t \left( P_t C_t + B_{t+1} - \left[ P_t W_t L_t + \Pi_t + R_{t-1} B_t - T_t + \Omega_t \right] \right) \right\} \right]$$

## Households: First-Order Conditions

• First-order condition for *C*<sub>t</sub>:

$$\frac{1}{C_t} = \mu_t P_t$$

• First-order condition for  $B_{t+1}$ :

 $\mu_t = \beta \mathbb{E}_t[\mu_{t+1}R_t]$ 

• Combine to consumption Euler equation (define  $\pi_{t+1} = P_{t+1}/P_t$ ):

$$\frac{1}{C_t} = \beta \mathbb{E}_t \left[ \frac{1}{C_{t+1}} \frac{R_t}{\pi_{t+1}} \right]$$

• Labor supply – first-order condition for *L<sub>t</sub>*:

$$\phi_t L_t^{\nu} = \mu_t P_t W_t = \frac{W_t}{C_t}.$$

- households;
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- exogenous shock processes

## **Final Goods Production**

• Production: (these guys just buy and combine intermediate goods)

$$Y_t = \left[\int_0^1 Y_t(i)^{\frac{1}{1+\lambda_t}} di\right]^{1+\lambda_t}$$

Profits

$$Y_tP_t - \int Y_t(i)P_t(i)di = \left[\int_0^1 Y_t(i)^{\frac{1}{1+\lambda_t}}di\right]^{1+\lambda_t}P_t - \int Y_t(i)P_t(i)di.$$

• Take prices as given and maximize profits by choosing optimal inputs  $Y_t(i)$ :

$$P_t(i) = P_t Y_t^{\lambda_t/(1+\lambda_t)} Y_t(i)^{-\lambda_t/(1+\lambda_t)} \implies Y_t(i) = \left(\frac{P_t(i)}{P_t}\right)^{-\frac{1+\lambda_t}{\lambda_t}} Y_t$$

• Free entry leads to zero profits:

$$Y_t P_t = \int Y_t(i) P_t(i) di \implies P_t = \left[\int_0^1 P_t(i)^{-\frac{1}{\lambda_t}} di\right]^{-\lambda_t}.$$

• Aggregate inflation is defined as  $\pi_t = P_t/P_{t-1}$ .

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### Intermediate Goods Production

• Production (these guys hire to produce something):

$$Y_t(i) = \max \left\{ A_t L_t(i) - \mathcal{F}, 0 \right\}.$$

• Firms are monopolistically competitive; face downward sloping demand curve:

$$Y_t(i) = \left(\frac{P_t(i)}{P_t}\right)^{-\frac{1+\lambda_t}{\lambda_t}} Y_t.$$

- Firms set prices to maximize profits, but there is a friction:
  - firms can only re-optimize their prices with probability  $1 \zeta_p$ ;
  - remaining  $1-\iota$  firms adjust their prices by  $ar{\pi}$
- Once prices are set, firms have to produce whatever quantity is demanded.

### Intermediate Goods Production

• Define the real marginal costs of producing a unit  $Y_{it}$  as

$$MC_t = \frac{W_t}{A_t}$$

• Decision problem ( $\beta^s \Xi_{t+s|t}$  is today's value of a future dollar)

$$\max_{\tilde{P}_{t}(i)} \qquad \mathbb{E}_{t} \left\{ \sum_{s=0}^{\infty} \zeta_{p}^{s} \beta^{s} \Xi_{t+s|t} Y_{t+s}(i) \left[ \tilde{P}_{t}(i) \bar{\pi}^{s} - P_{t+s} M C_{t+s} \right] \right\}$$
  
s.t. 
$$Y_{t+s}(i) = \left( \frac{\tilde{P}_{t}(i) \bar{\pi}^{s}}{P_{t+s}} \right)^{-\frac{1+\lambda_{t}}{\lambda_{t}}} Y_{t+s}$$

• Differentiate with respect to  $\tilde{P}_t(i)$  to obtain first-order condition for optimal price.

- households;
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- We did not specify a money demand equation, but we could. It would depend on the nominal interest rate. The higher *R*<sub>t</sub>, the lower the demand for money.
- Central bank prints enough money so that demand is satisfied at interest rate implied by monetary policy rule:

$$R_{t} = R_{*,t}^{1-\rho_{R}} R_{t-1}^{\rho_{R}} \exp\{\sigma_{R} \epsilon_{R,t}\}, \quad R_{*,t} = (r\pi_{*}) \left(\frac{\pi_{t}}{\pi_{*}}\right)^{\psi_{1}} \left(\frac{Y_{t}}{\gamma Y_{t-1}}\right)^{\psi_{2}}$$

- r is equilibrium real rate.
- $\pi_*$  is target inflation rate.
- $\epsilon_{R,t}$  is exogenous monetary policy shock. Interpretation?

- For now, it's passive and not very interesting.
- Budget constraint:

$$P_t G_t + R_{t-1} B_t + M_t = T_t + B_t + M_{t+1}$$

- Lump-sum taxes/transfer balance the budget in every period. Seigniorage does not matter.
- Government spending is exogenous. Re-scale:

$$G_t = \left(1 - rac{1}{g_t}
ight) Y_t.$$

- households;
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## Exogenous shock processes

- Total factor productivity A<sub>t</sub>.
- Preference / labor demand shifter  $\phi_t$ .
- Mark-up shock  $\lambda_t$ .
- Monetary policy shock  $\epsilon_{R,t}$ .
- Government spending shock  $g_t$ .
- We will specify exogenous laws of motions for these processes, e.g.,

$$\ln g_t = (1 - \rho_g) \ln g^* + \rho_g \ln g_{t-1} + \sigma_g \epsilon_{g,t}, \quad \epsilon_{g,t} \sim N(0,1).$$

## A Small-Scale New Keynesian DSGE Model

So far:

- households;
- final goods producing firms;
- intermediate goods producing firms;
- central bank and fiscal authority;
- exogenous shock processes.

#### What's left to do?

- derive aggregate resource constraint;
- derive evolution of price dispersion;
- complete markets assumption;

- After deriving the equilibrium conditions of the model, we now need to solve for the dynamics of the endogenous variables.
- System of nonlinear expectational difference equations;
- Find solution(s) of system of expectational difference equations:
  - global (nonlinear) approximation methods;
  - local approximation near steady state.
- We will focus on log-linear approximations around the steady state.
- Many more details in FVRRS.

# Our Goal: State-space Representation of DSGE Model

•  $n_{y} \times 1$  vector of observables:

$$y_t = M_y' [\log(X_t/X_{t-1}), \log lsh_t, \log \pi_t, \log R_t]'.$$

•  $n_s imes 1$  vector of econometric state variables  $s_t$ 

$$s_t = [\phi_t, \lambda_t, z_t, \epsilon_{R,t}, \widehat{x}_{t-1}]'$$

• DSGE model parameters:

$$\theta = [\beta, \gamma, \lambda, \pi^*, \zeta_{\mathbf{p}}, \nu, \rho_{\phi}, \rho_{\lambda}, \rho_z, \sigma_{\phi}, \sigma_{\lambda}, \sigma_z, \sigma_R]'.$$

• Measurement equation:

$$y_t = \Psi_0(\theta) + \Psi_1(\theta)s_t.$$

• State-transition equation:

$$s_t = \Phi_1(\theta)s_{t-1} + \Phi_\epsilon(\theta)\epsilon_t, \quad \epsilon_t = [\epsilon_{\phi,t}, \epsilon_{\lambda,t}, \epsilon_{z,t}, \epsilon_{R,t}]'$$

#### Our Goal: State-Space Representation of DSGE Model

State-space representation:

 $y_t = \Psi_0(\theta) + \Psi_1(\theta)s_t$  $s_t = \Phi_1(\theta)s_{t-1} + \Phi_\epsilon(\theta)\epsilon_t$ 

System matrices:

$$\begin{split} \Psi_{0}(\theta) &= M_{y}' \begin{bmatrix} \log(s) \\ \log(s) \\ \log(s) \\ \log(\pi^{*}) \\ \log(\pi^{*}\gamma/\beta) \end{bmatrix}, \quad x_{\phi} = -\frac{\kappa_{\rho}\psi_{\rho}/\beta}{1 - \psi_{\rho}\rho_{\phi}}, \quad x_{\lambda} = -\frac{\kappa_{\rho}\psi_{\rho}/\beta}{1 - \psi_{\rho}\rho_{\lambda}}, \quad x_{z} = \frac{\rho_{z}\psi_{p}}{1 - \psi_{\rho}\rho_{z}}, \quad x_{e_{R}} = -\psi_{\rho}\sigma_{R} \end{split}$$

$$\begin{split} \Psi_{1}(\theta) &= M_{y}' \begin{bmatrix} x_{\phi} & x_{\lambda} & x_{z}+1 & x_{e_{R}} & -1 \\ 1 + (1 + \nu)x_{\phi} & (1 + \nu)x_{\lambda} & (1 + \nu)x_{z} & (1 + \nu)x_{e_{R}} & 0 \\ \frac{\kappa_{\rho}}{1 - \beta\rho_{\phi}}(1 + (1 + \nu)x_{\phi}) & \frac{\kappa_{\rho}}{1 - \beta\rho_{\lambda}}(1 + (1 + \nu)x_{\lambda}) & \frac{\kappa_{\rho}}{1 - \beta\rho_{z}}(1 + \nu)x_{z} & +\kappa_{\rho}(1 + \nu)x_{e_{R}} & 0 \\ \frac{\kappa_{\rho}/\beta}{1 - \beta\rho\phi}(1 + (1 + \nu)x_{\phi}) & \frac{\kappa_{\rho}/\beta}{1 - \beta\rho_{\lambda}}(1 + (1 + \nu)x_{\lambda}) & \frac{\kappa_{\rho}/\beta}{1 - \beta\rho_{z}}(1 + \nu)x_{z} & (\kappa_{\rho}(1 + \nu)x_{e_{R}}/\beta + \sigma_{R}) & 0 \end{bmatrix} \end{split}$$

$$\begin{split} \Phi_{1}(\theta) &= \begin{bmatrix} \rho_{\phi} & 0 & 0 & 0 \\ 0 & \rho_{\lambda} & 0 & 0 \\ 0 & 0 & \rho_{z} & 0 \\ 0 & 0 & 0 & 0 \\ \kappa_{\phi} & x_{\lambda} & x_{z} & x_{e_{R}} \end{bmatrix}, \quad \Phi_{e}(\theta) = \begin{bmatrix} \sigma_{\phi} & 0 & 0 & 0 \\ 0 & \sigma_{\lambda} & 0 & 0 \\ 0 & 0 & \sigma_{z} & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}$$

 ${\it M}_y'$  is an  ${\it n}_y$   $\times$  4 selection matrix that selects rows of  $\Psi_0$  and  $\Psi_1.$ 

# Steady State

Shut down aggregate uncertainty: set all shock standard deviations  $\sigma_{\cdot} = 0$ .

• Technology:

$$\ln A_t = \ln \gamma + \ln A_{t-1} + z_t, \quad z_t = \rho_z z_{t-1} + \sigma_z \epsilon_{z,t}.$$

Set  $\sigma_z = 0$ :  $\ln A_t^* = \gamma t$ .

• Preferences:

$$\ln \phi_t = (1 - \rho_\phi) \ln \phi + \rho_\phi \ln \phi_{t-1} + \sigma_\phi \epsilon_{\phi,t}.$$

• Mark-up:

$$\ln \lambda_t = (1 - \rho_\lambda) \ln \lambda + \rho_\lambda \ln \lambda_{t-1} + \sigma_\lambda \epsilon_{\lambda,t}.$$

#### • Government Spending:

$$\ln g_t = (1 - \rho_g) \ln g^* + \rho_g \ln g_{t-1} + \sigma_g \epsilon_{g,t}$$

- Problem: this economy grows... which does not lead to a steady state.
- Solution: detrend model variables by  $A_t$ .
- Model has steady state in terms of detrended variables.

### Example: Detrend Households' Euler Equation

• Recall:

$$\frac{1}{C_t} = \beta \mathbb{E}_t \left[ \frac{1}{C_{t+1}} \frac{R_t}{\pi_{t+1}} \right]$$

• Rewrite:

$$\frac{A_t}{C_t} = \beta \mathbb{E}_t \left[ \frac{A_{t+1}}{C_{t+1}} \frac{A_t}{A_{t+1}} \frac{R_t}{\pi_{t+1}} \right] \quad \Longrightarrow \quad \frac{1}{c_t} = \beta \mathbb{E}_t \left[ \frac{1}{c_{t+1}} \frac{1}{\gamma e^{z_{t+1}}} \frac{R_t}{\pi_{t+1}} \right]$$

• Steady state:

$$R = \pi \frac{\gamma}{\beta} = \pi r.$$

## Local Approximation Around Steady State

- We will now approximate the equilibrium dynamics of the model.
- Taylor series expansion around around the steady state.

# What is a Log-Linear Approximation?

• Consider Cobb-Douglas production function:  $Y_t = Z_t K_t^{\alpha} H_t^{1-\alpha}$ .

• Linearization around 
$$Y_*$$
,  $Z_*$ ,  $K_*$ ,  $H_*$ :  
 $Y_t - Y_* = K_*^{\alpha} H_*^{1-\alpha} (Z_t - Z_*) + \alpha Z_* K_*^{\alpha-1} H_*^{1-\alpha} (K_t - K_*)$   
 $+ (1 - \alpha) Z_* K_*^{\alpha} H_*^{-\alpha} (H_t - K_*)$ 

• Log-linearization: Let  $f(x) = f(e^{v})$  and linearize with respect to v:

$$f(e^{v}) pprox f(e^{v_*}) + e^{v_*} f'(e^{v_*})(v-v_*).$$

Thus:

$$f(x) \approx f(x_*) + x_* f'(x_*) (\ln x/x_*) = f(x_*) + f'(x_*) \hat{x}$$

• Cobb-Douglas production function:

$$\tilde{Y}_t = \hat{Z}_t + \alpha \hat{K}_t + (1 - \alpha) \hat{H}_t$$

## Let's Try the Log-linearizations

• Euler Equation:

$$\frac{1}{c_t} = \beta \mathbb{E}_t \left[ \frac{1}{c_{t+1}} \frac{1}{\gamma e^{z_{t+1}}} \frac{R_t}{\pi_{t+1}} \right].$$

• Log-linearized:

$$-\widehat{c}_t = \mathbb{E}_t \Big[ -\widehat{c}_{t+1} - z_{t+1} + \widehat{R}_t - \widehat{\pi}_{t+1} \Big] \implies \widehat{c}_t = \mathbb{E}_t [\widehat{c}_{t+1}] - (\widehat{R}_t - \mathbb{E}[\widehat{\pi}_{t+1}]) + \mathbb{E}_t [z_{t+1}].$$

• Labor Supply:

$$\phi_t L_t^{\nu} = \frac{w_t}{c_t}.$$

• Log-linearized:

$$\widehat{\phi}_t + \nu \widehat{L}_t = \widehat{w}_t - \widehat{c}_t$$

# Combining Bits and Pieces

• Notation: write  $x_t$  instead of  $y_t$  for output.

• Assume: 
$$\pi = \bar{\pi} = \pi_*$$
,  $\psi_1 = 1/\beta$ ,  $\psi_2 = 0$ ,  $\rho_R = 0$ .

• Linear rational expectations (LRE) system:

$$\begin{aligned} \widehat{c}_t &= \mathbb{E}_{t+1}[\widehat{c}_{t+1}] - \left(\widehat{R}_t - \mathbb{E}_t[\widehat{\pi}_{t+1}]\right) + \mathbb{E}_t[z_{t+1}] \\ \widehat{\pi}_t &= \beta \mathbb{E}_t[\widehat{\pi}_{t+1}] + \kappa_\rho(\widehat{lsh}_t + \lambda_t) \\ \widehat{R}_t &= \frac{1}{\beta}\widehat{\pi}_t + \sigma_R \epsilon_{R,t} \\ \widehat{lsh}_t &= (1+\nu)\widehat{c}_t + \nu \widehat{g}_t + \phi_t \\ \widehat{x}_t &= \widehat{c}_t + \widehat{g}_t \\ \widehat{g}_t &= \rho_g \widehat{g}_{t-1} + \sigma_g \epsilon_{g,t} \\ \phi_t &= \rho_\phi \phi_{t-1} + \sigma_\phi \epsilon_{\phi,t} \\ \lambda_t &= \rho_\lambda \lambda_{t-1} + \sigma_\lambda \epsilon_{\lambda,t} \\ z_t &= \rho_z z_{t-1} + \sigma_z \epsilon_{z,t} \end{aligned}$$

# How Can One Solve LRE Systems? A Simple Example

Simple model:

$$y_t = rac{1}{ heta} \mathbb{E}_t[y_{t+1}] + \epsilon_t, \quad \epsilon_t \sim \textit{iid}(0,1), \quad heta \in \Theta = [0,2].$$

**Sims (2002) Method:** Introduce conditional expectation  $\xi_t = \mathbb{E}_t[y_{t+1}]$  and forecast error  $\eta_t = y_t - \xi_{t-1}$ :

$$\xi_t = \theta \xi_{t-1} - \theta \epsilon_t + \theta \eta_t.$$

Nonexplosive solutions:

• Determinacy:  $\theta > 1$ . The only stable solution:

$$\xi_t = 0, \quad \eta_t = \epsilon_t \implies y_t = \epsilon_t$$

• Indeterminacy:  $\theta \leq 1$  the stability requirement imposes no restrictions on forecast error:

$$\eta_t = \widetilde{M} \epsilon_t + \frac{\zeta_t}{\zeta_t} \implies y_t = \theta y_{t-1} + \widetilde{M} \epsilon_t + \frac{\zeta_t}{\zeta_t} - \theta \epsilon_{t-1}$$

- A simplified version of our DSGE model can be solved "by hand."
- More realistic models need to be solved numerically.
- The numerical solution can be expressed as a VAR(1) in terms of suitably chosen model variables *s*<sub>t</sub>:

 $s_t = \Phi_1( heta) s_{t-1} + \Phi_\epsilon( heta) \epsilon_t.$ 

 Many solution techniques are available for LRE models, e.g., Blanchard and Kahn (1980), King and Watson (1998), Uhlig (1999), Anderson (2000), Klein (2000), Christiano (2002).

- To confront the model with data, one has to account for the presence of the model-implied stochastic trend in aggregate output and to add the steady states to all model variables.
- Measurement equations:

$$\begin{split} \log(X_t/X_{t-1}) &= \widehat{x}_t - \widehat{x}_{t-1} + z_t + \log \gamma \\ \log(lsh_t) &= \widehat{lsh}_t + \log(lsh) \\ \log \pi_t &= \widehat{\pi}_t + \log \pi^* \\ \log R_t &= \widehat{R}_t + \log(\pi^*\gamma/\beta). \end{split}$$

# State-space Representation of DSGE Model

•  $n_{\gamma} \times 1$  vector of observables:

$$y_t = M_y' [\log(X_t/X_{t-1}), \log lsh_t, \log \pi_t, \log R_t]'.$$

•  $n_s imes 1$  vector of econometric state variables  $s_t$ 

$$s_t = [\phi_t, \lambda_t, z_t, \epsilon_{R,t}, \widehat{x}_{t-1}]'$$

• DSGE model parameters:

$$\theta = [\beta, \gamma, \lambda, \pi^*, \zeta_{\mathbf{p}}, \nu, \rho_{\phi}, \rho_{\lambda}, \rho_z, \sigma_{\phi}, \sigma_{\lambda}, \sigma_z, \sigma_{\mathbf{R}}]'.$$

• Measurement equation:

$$y_t = \Psi_0(\theta) + \Psi_1(\theta)s_t.$$

• State-transition equation:

$$s_t = \Phi_1(\theta)s_{t-1} + \Phi_\epsilon(\theta)\epsilon_t, \quad \epsilon_t = [\epsilon_{\phi,t}, \epsilon_{\lambda,t}, \epsilon_{z,t}, \epsilon_{R,t}]'$$

### State-Space Representation of DSGE Model

State-space representation:

 $y_t = \Psi_0(\theta) + \Psi_1(\theta)s_t$  $s_t = \Phi_1(\theta)s_{t-1} + \Phi_\epsilon(\theta)\epsilon_t$ 

System matrices:

 ${\it M}_y'$  is an  ${\it n}_y$   $\times$  4 selection matrix that selects rows of  $\Psi_0$  and  $\Psi_1.$ 

- Simulation: pick values for parameter vector  $\theta \implies$  simulate data  $Y^{sim}(\theta)$  and determine its properties.
- Statistical inference:
  - observed data  $Y^{obs} \Longrightarrow$  determine suitable values for parameter vector  $\theta$ .
  - **Basic Idea:** choose  $\theta$  such that  $Y^{sim}(\theta)$  look like  $Y^{obs}$ .
  - Goals: estimates  $\hat{\theta}$  as well as measures of uncertainty associated with these estimates.

#### • Frequentist:

- pre-experimental perspective;
- condition on "true" but unknown  $\theta_0$ ;
- treat data Y as random;
- study behavior of estimators and decision rules under repeated sampling.

#### • Bayesian:

- post-experimental perspective;
- condition on observed sample Y;
- treat parameter  $\theta$  as unknown and random;
- derive estimators and decision rules that minimize expected loss (averaging over  $\theta$ ) conditional on observed Y.

**DSGE model**  $(M_1)$  is assumed to be correctly specified, i.e. we believe the probabilistic structure is rich enough to assign high probability to the salient features of macroeconomic time series.

- Desirable to let the model-implied probability distribution  $p(Y|\theta_0, M_1)$  determine the choice of the objective function for estimators and test statistics to obtain a statistical procedure that is efficient (meaning that the estimator is close to  $\theta_0$  with high probability in repeated sampling).
- Maximum likelihood (ML) estimator

 $\hat{\theta}_{ml} = \operatorname{argmax}_{\theta \in \Theta} \log p(Y|\theta, M_1).$ 

• Minimize discrepancy between sample statistics  $\hat{m}_{\mathcal{T}}(Y)$  and model-implied population statistics  $\mathbb{E}[\hat{m}_{\mathcal{T}}(Y)|\theta, M_1]$ :

$$\hat{ heta}_{md} = \operatorname{argmin}_{ heta \in \Theta} \, Q_{\mathcal{T}}( heta|Y) = \left\| \hat{m}_{\mathcal{T}}(Y) - \mathbb{E}[\hat{m}_{\mathcal{T}}(Y)| heta, M_1] 
ight\|_{W_{\mathcal{T}}},$$

**DSGE model**  $(M_1)$  is assumed to be correctly specified, i.e. we believe the probabilistic structure is rich enough to assign high probability to the salient features of macroeconomic time series.

- Initial state of knowledge summarized in **prior** distribution  $p(\theta)$ .
- Update in view of data Y to obtain **posterior** distribution  $p(\theta|Y)$ :

$$p(\theta|Y,M_1) = \frac{p(Y|\theta,M_1)p(\theta|M_1)}{p(Y|M_1)}, \quad p(Y|M_1) = \int p(Y|\theta,M_1)p(\theta|M_1)d\theta.$$

• Make decisions that minimize posterior expected loss:

$$\delta_* = \operatorname{argmin}_{\delta \in \mathcal{D}} \int L(h(\theta), \delta) p(\theta | Y, M_1) d\theta.$$

• Place probabilities on competing models and update:

$$\frac{\pi_{1,T}}{\pi_{2,T}} = \frac{\pi_{1,0}}{\pi_{2,0}} \frac{p(Y|M_1)}{p(Y|M_2)}.$$

## Model Misspecification is a Concern



### Frequentist Inference

DSGE model  $(M_1)$  is assumed to be misspecified or incompletely specified.

• Example: suppose our DSGE model only has a monetary policy shock. Then,

$$\frac{1}{\kappa_p(1+\nu)x_{\epsilon_R}/\beta+\sigma_R}\widehat{R}_t-\frac{1}{\kappa_p(1+\nu)x_{\epsilon_R}}\widehat{\pi}_t=0,$$

which is clearly violated in the data.

- Need reference model  $M_0$ , e.g., VAR, under which to evaluate sampling distribution of Y.
- Concept of "true" value is no longer sensible  $\implies$  pseudo-optimal parameter value:

 $\theta_0(Q, W) = \operatorname{argmin}_{\theta \in \Theta} Q(\theta | M_0),$ 

where

$$Q(\theta|M_0) = \left\| \mathbb{E}[\hat{m}_T(Y)|M_0] - \mathbb{E}[\hat{m}(Y)|\theta, M_1] \right\|_W.$$

# **Bayesian Inference**

#### DSGE model $(M_1)$ is assumed to be misspecified or incompletely specified.

- Derive posterior distributions under a more flexible reference model  $M_0$ , e.g., VAR. Then choose  $\theta$  to minimize discrepancy between implications of  $M_0$  and DSGE model  $M_1$ .
- Use DSGE model  $M_1$  to generate a prior distribution for a more flexible reference model  $M_0$ .
- Rather than using posterior probabilities to select among or average across two DSGE models, one can form a prediction pool, which is essentially a linear combination of two predictive densities:

$$\lambda p(y_t|Y_{1:t-1}, M_1) + (1-\lambda)p(y_t|Y_{1:t-1}, M_2).$$

The weight  $\lambda \in [0,1]$  can be determined based on

$$\prod_{t=1}^{T} \left[ \lambda p(y_t | Y_{1:t-1}, M_1) + (1-\lambda) p(y_t | Y_{1:t-1}, M_2) \right]$$