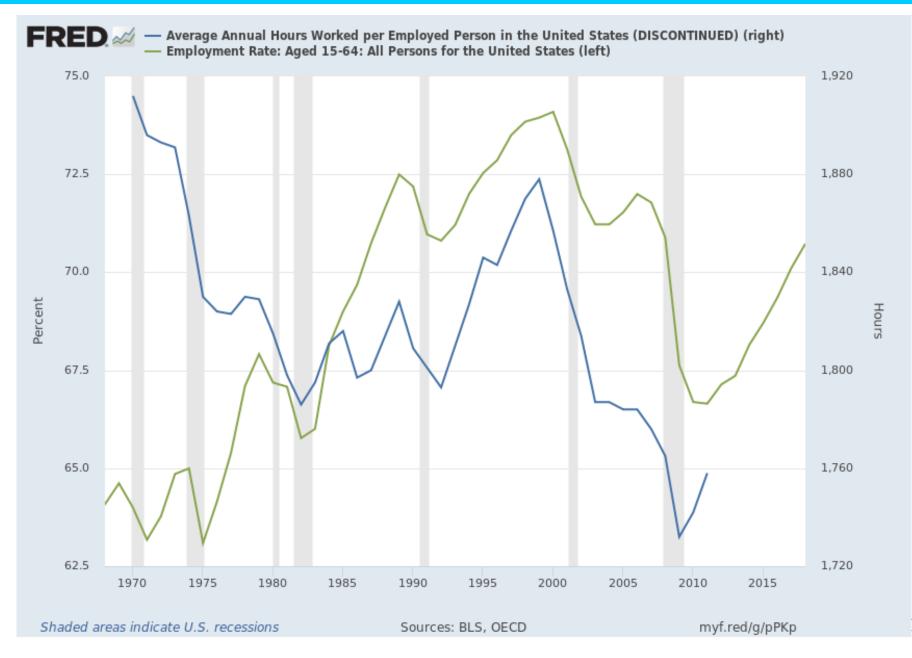
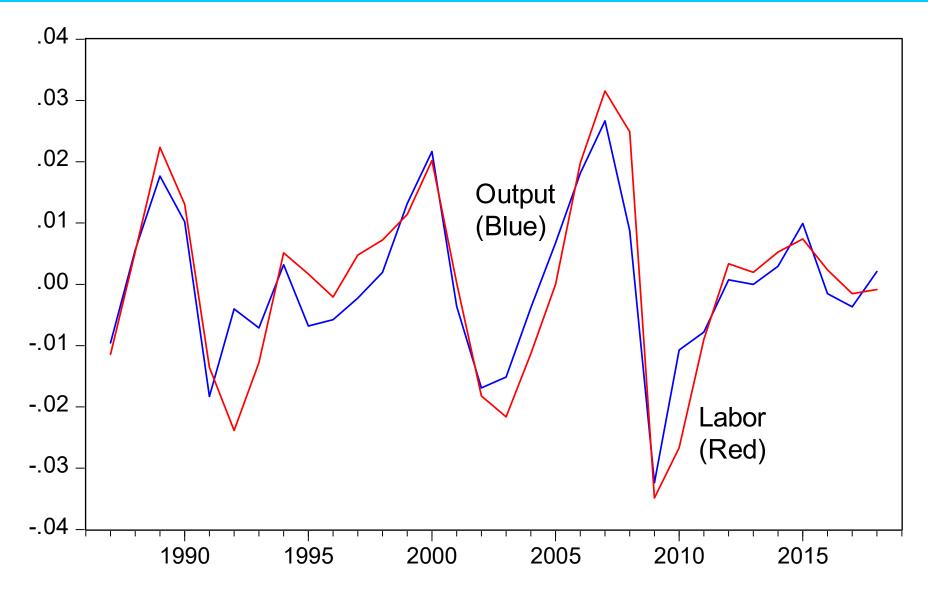
# **Overview and Labor Wedge**

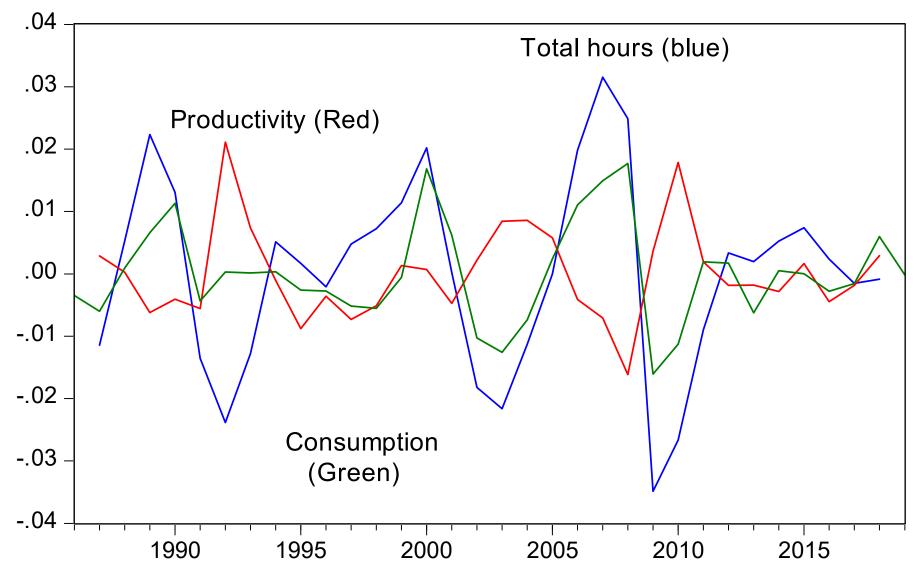
# **U.S. Emp/Pop and Hours per Worker**



#### **Cyclicality: Real Output and Total Hours** (Private sector, HP filtered)



#### **Cyclicality: Labor Productivity and Consumption** (HP filtered, Cons = Nondurs and Services)

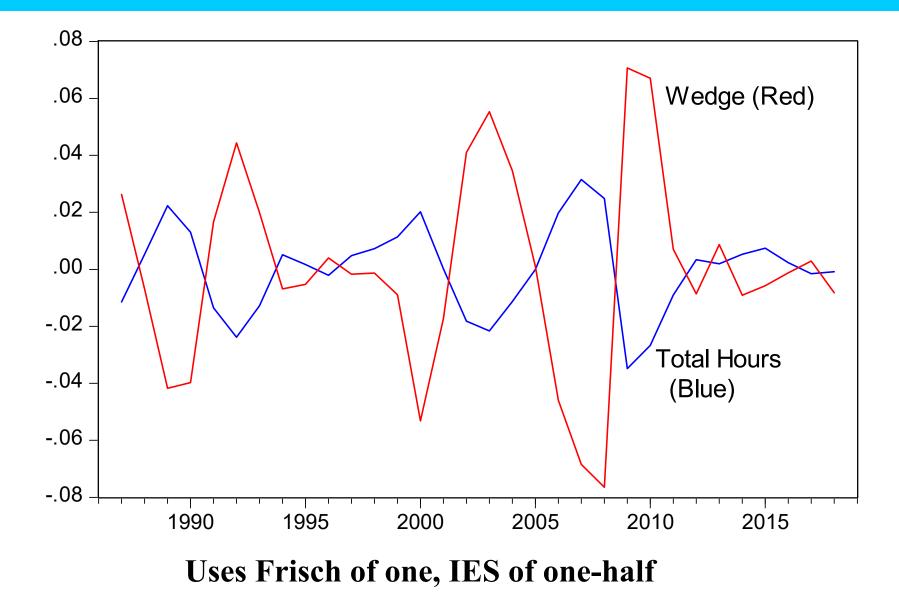


# U.S. Labor Wedge, 1987 to 2018

	Elasticity with respect to:		
	Real GDP	Total Hours	
Labor Productivity	- 0.26 (.10)	- 0.33 (.08)	
Total hours	1.48 (.10)	1	
Consumption	0.71 (08)	0.43 (.06)	
Wedge	- 3.16 (.31)	-2.19 (.11)	

Notes: Total hours and labor productivity is for private economy; GDP includes government sector. Consumption is nondurables and services. Sample covers 1987 to 2018. All series are logged and HP-filtered. The wedge assumes an IES of 0.5 and a Frisch of 1.0.

#### **Cyclicality in Wedge versus Total Hours**



# **Labor Supply**

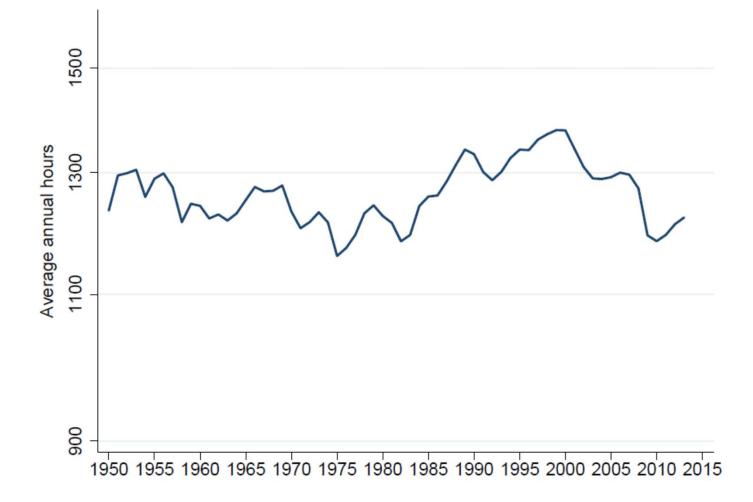
# **Boppart and Krusell Overview**

- Standard Macro model assumes balanced growth path, with constant hours worked
  - But data show declining hours (will see figures)
  - Looks roughly like linear trend (constant negative growth rate) in Ln(hours)
- Is also consistent with higher hours worked in poorer countries (Bick et. al., will show below)
- Consider preferences that produce balanced growth with declining hours: requires stronger wealth effect on leisure than in KPR

#### **Intuition for Preferences**

In compact terms, one can describe the period utility function under KPR as a power function of cv(h), where c is consumption and h hours worked and v is an arbitrary (decreasing) function. What we show in our main Theorem 1 is that the broader class has the same form: period utility is a power function of  $cv(hc^{\frac{\nu}{1-\nu}})$ , where  $\nu < 1$  is the preference parameter that guides how fast hours shrink relative to productivity. In terms of gross rates, if productivity grows at rate  $\gamma$ , then hours grow at rate  $\gamma^{-\nu}$ , whereas consumption grows at  $\gamma^{1-\nu}$ . For  $\nu > 0$ , the factor  $c^{\frac{\nu}{1-\nu}}$  captures the stronger income effect: as consumption grows, there is an added "penalty" to working (since v is decreasing). Our preference class obviously nests KPR: KPR corresponds to  $\nu = 0$ .

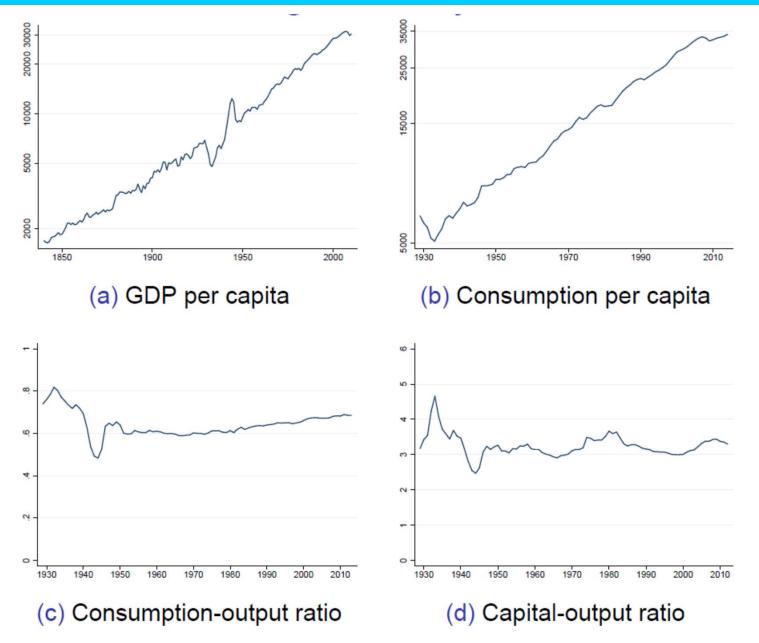
#### Little trend in U.S. hours post WWII



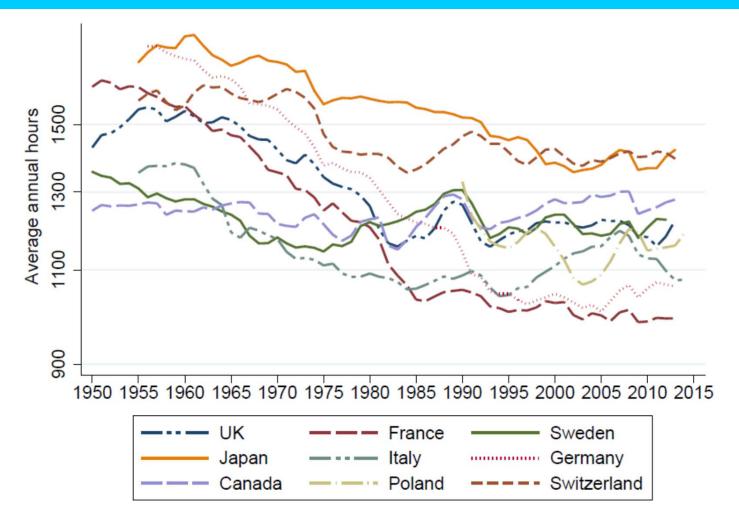
#### Figure: U.S. average annual hours per capita aged 15-64, 1950-2013

**Notes:** Source: GGDC Total Economy Database for total hours worked and OECD for the data on population aged 15–64. The figure is comparable to the ones in Rogerson (2006). Regressing the logarithm of hours worked on time gives an insignificant slope coefficient.

#### **U.S. balanced growth stylized facts**



#### **But decline in many countries**

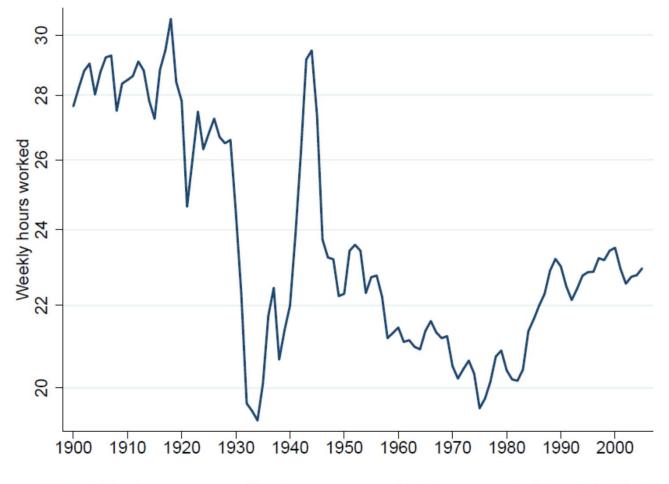


# Figure: Selected countries average annual hours per capita aged 15–64, 1950–2015

Notes: Source: GGDC Total Economy Database for total hours worked and OECD for the data on population aged 15–64. The figure is comparable to the ones in Rogerson (2006). Regressing the logarithm of hours worked on time gives a slope coefficient of -0.00455. 1.12

#### **Declined historically in U.S.**

U.S. data including the pre-war period



#### Figure: Weekly hours worked per population aged 14+, 1900–2005

Notes: Source: Ramey and Francis (2009). Regressing the logarithm of hours worked on time gives slope coefficient of -0.00285.

# At intensive margin

#### Intensive and extensive margin over 100+ years



Figure: Hours per worker and participation rate in the U.S.

**Notes:** The scale is logarithmic in the figure on hours worked per worker. Regressing the logarithm of hours worked per worker on time gives slope coefficient of -0.00418. Source: Ramey and Francis (2009).

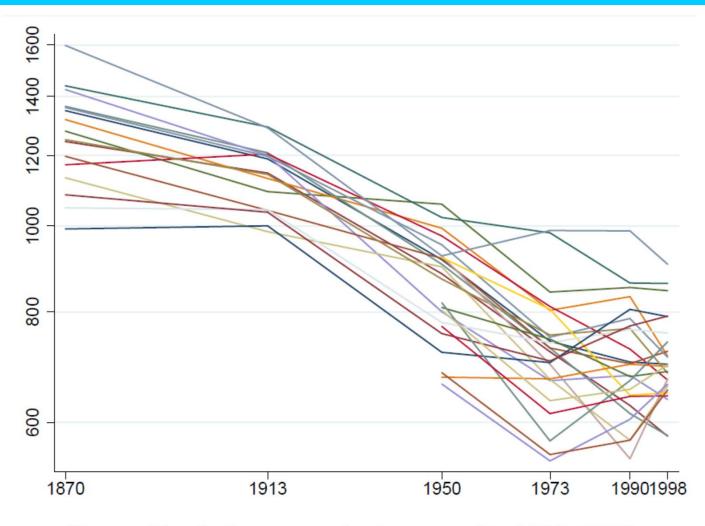
#### U.S. workweek back to 1830



#### Figure: U.S. weekly hours worked in nonfarm establishments 1830–2015

**Source:** Average weekly hours data for 1830–80: Whaples (1990, Table 2.1). 1890-1970: Historical Statistics of the United States: Colonial Times to 1970 (Series D765 and D803). 1970–2015: Statistical Abstract of the United States the number for nonfarm establishments. This graph shows an updates series of the data in Greenwood and Vandenbroucke (2008). Regressing the log of hours on a constant and year gives a slope coefficient of -0.00315 in the full sample (and -0.00208 for the years 1970–2015).

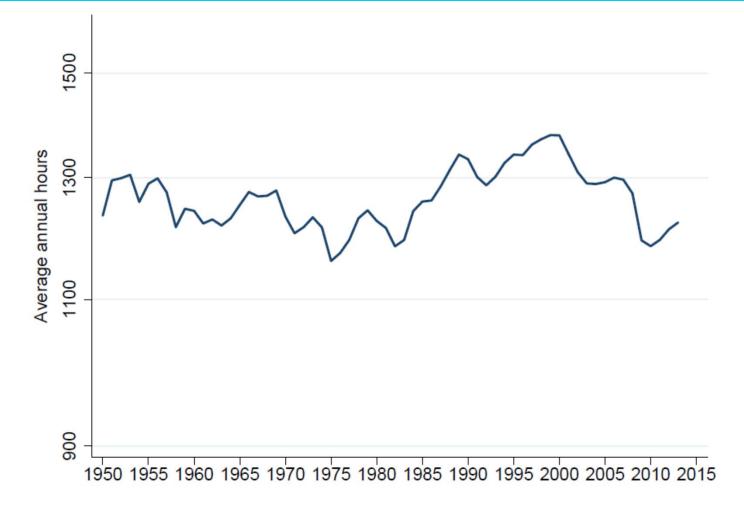
#### **Declined historically elsewhere**



#### Figure: Yearly hours worked per capita 1870–1998

**Source:** Maddison (2001). The sample includes the following 25 countries: Austria, Belgium, Denmark, Finland, France, Germany, Italy, Netherlands, Norway, Sweden, Switzerland, United Kingdom, Ireland, Spain, Australia, Canada, United States, Argentina, Brazil, Chile, Colombia, Mexico, Peru, Venezuela, Japan. Regressing the log of hours on a country fixed effect and year gives a slope coefficient of -0.00462 in the full sample (and -0.00398 for the period 1950–1998).

# So post WWII not representative



#### Figure: U.S. average annual hours per capita aged 15-64, 1950-2013

**Notes:** Source: GGDC Total Economy Database for total hours worked and OECD for the data on population aged 15–64. The figure is comparable to the ones in Rogerson (2006). Regressing the logarithm of hours worked on time gives an insignificant slope coefficient.

#### **Important caveats**

• Leisure has notably increased (Aguiar & Hurst)

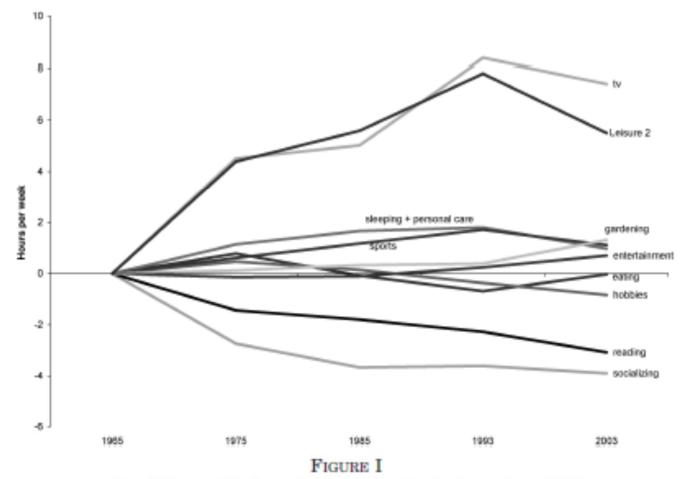
#### Leisure since 1965, Aguiar & Hurst

HOURS PER WEEK SPENT IN LEISURE FOR FULL SAMPLE, MEN, AND WOMEN

	Average hours per week spent in leisure					
Time-use category (hours per week)	1965	1975	1985	1993	2003	Difference: 2003–1965
Panel 1: Full sample						
Leisure Measure 1	30.77	33.24	34.78	37.47	35.33	4.56
Leisure Measure 2	102.23	106.62	107.82	110.04	107.73	5.50
Leisure Measure 3	105.90	109.74	111.46	113.16	113.23	7.33
Leisure Measure 4	109.93	114.06	114.33	116.39	117.98	8.05
Panel 2: Men						
Leisure Measure 1	31.80	33.36	35.15	37.65	37.40	5.60
Leisure Measure 2	101.68	105.33	106.81	108.50	107.88	6.20
Leisure Measure 3	103.12	106.73	108.47	109.97	111.13	8.01
Leisure Measure 4	106.75	110.62	110.68	112.82	115.04	8.29
Panel 3: Women						
Leisure Measure 1	29.89	33.14	34.46	37.32	33.54	3.65
Leisure Measure 2	102.70	107.75	108.69	111.38	107.59	4.89
Leisure Measure 3	108.31	112.35	114.05	115.92	115.06	6.75
Leisure Measure 4	112.69	117.05	117.49	119.48	120.52	7.83

All means are calculated using fixed demographic weights, as described in the text. Leisure Measure 1 refers to the time individuals spent socializing, in passive leisure, in active leisure, volunteering, in pet care, and gardening. Leisure Measure 2 refers to the time individuals spent in Leisure Measure 1 plus time spent sleeping, eating, and in personal activities (excluding own medical care). Leisure Measure 3 includes Leisure Measure 2 plus time spent in child care. Leisure Measure 4 is defined as any time not allocated to market or nonmarket work. See Table IX and text for additional detail. The relevant sample sizes are as reported in Table II. The sample restrictions are described in the footnote to Table I.

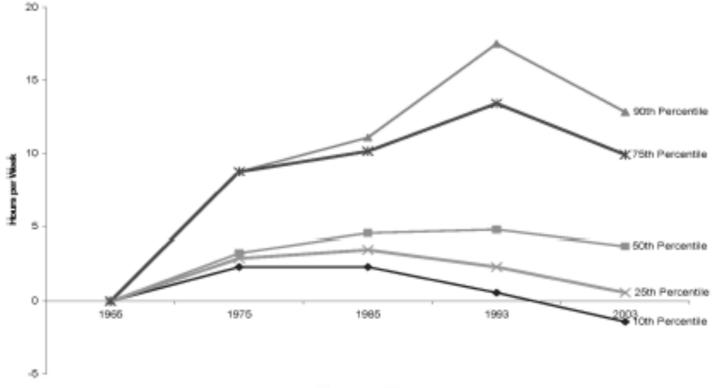
#### Leisure since 1965, Aguiar & Hurst





This figure plots the evolution of the subcomponents of Leisure 2 for the full sample, represented as differences from each subcomponent's mean in 1965. All means are calculated using fixed demographic weights, as described in the text.

### **Rise in leisure inequality**

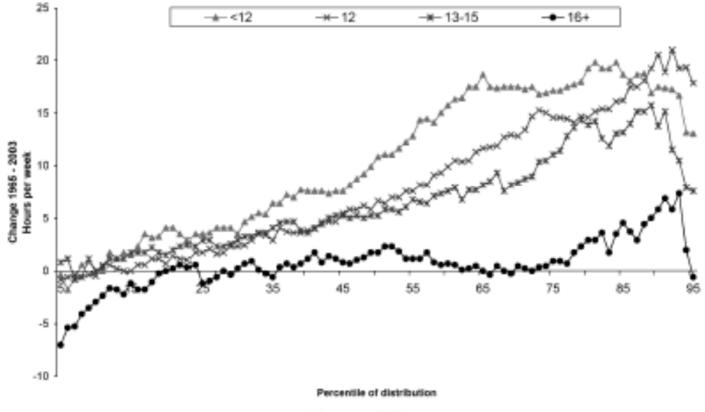


#### FIGURE II

Key Percentiles of Leisure 2 Distribution, Deviations from 1965

This figure plots the evolution of key percentiles of the cross-sectional distribution of Leisure 2 for the full sample, represented as differences from each percentile point's value in 1965. The percentile points represent the unconditional sample distribution in each year, unadjusted for demographic changes.

#### Leisure shifted to lower-waged workers



#### FIGURE IV

Change by Percentile Point for Leisure 2 by Educational Attainment 1965–2003

This figure plots the change at each percentile point of the Leisure 2 distribution between 1965 and 2003, broken down by educational attainment. The percentile points represent the unconditional distribution of the respective subsample in each year, unadjusted for demographic changes.

#### **Important caveats**

- Leisure has notably increased (Aguiar & Hurst)
- Models abstract from a number of factors
  - Nature of work/leisure dramatically evolves
  - Innovations in home production—increased market labor
  - The variety of market goods evolves: encouraged market labor
  - Nature of leisure activities evolved—ambiguous effect

#### Bick, et al., "How Do Hours worked vary with income?"

- Compare employment rates and average hours across large set of countries—compared as of year 2000, for ages 25-54
- Focus on 48 core countries with similar samples/definitions
  - Workers working in sectors measured in GDP: includes agric. & self-employed, but not home sector
  - Respondents report actual hours worked over recent time period (last week, month); data collected over entire calendar year

#### **Bick, et al., Main findings**

- Rich countries (top 3<sup>rd</sup> GDP) work 18.9 hours per week, compared to 28.5 in poor countries (bottom 3<sup>rd</sup> GDP), 40 percent difference in logs
  - Elasticity of hours wrt GDP/hour is -0.12
  - Employ. rates account for 3/4ths (same as for business cycles)
  - Expands welfare differences, about 40% in income units—high income/low factor 19 rather than 12 (nature of work also different)
- Within countries
  - Relative hours fall with relative wage in most countries, but in richest countries do not (in poorer countries fit relation across countries)
  - Is stronger for men--reflect lack of non-market info?

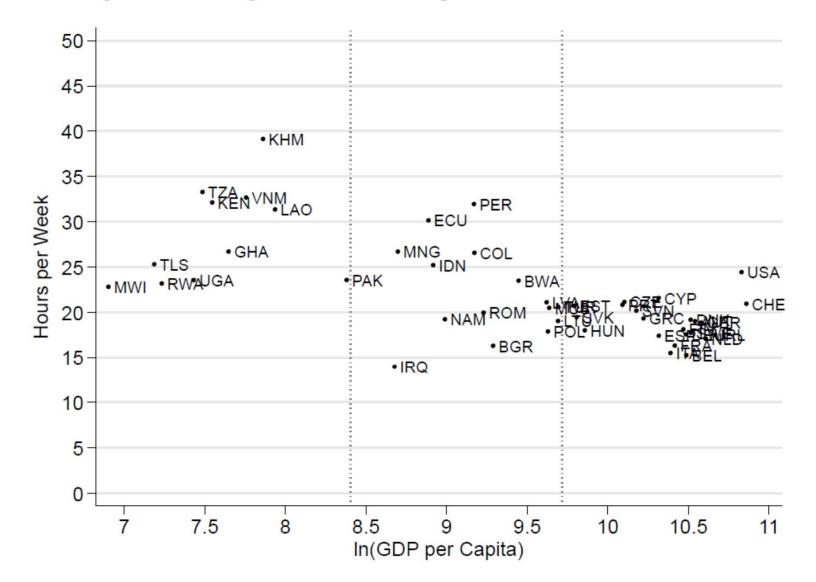
# **Main cross-section**

Table 2: Employment Rates and Hours Per Employed

	С	ountry Income Gro	up
	Low	Middle	High
Hours Per Adult	28.5	22.2	18.9
Employment Rate	75.3	53.7	54.9
Hours Per Worker	38.4	41.2	34.5

#### Main cross-section cont.

Figure 1: Average Hours Worked per Adult in Core Countries



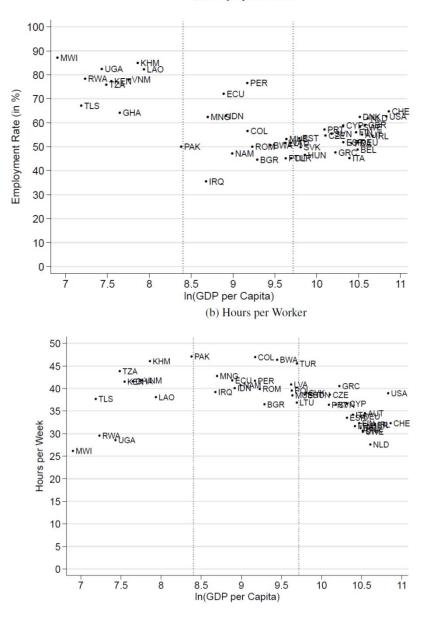
1.27

# Holds controlling for gender, education

Sex	Country	Country Income Group			
	Low M	Middle			
All	28.5	22.2 18			
Women	24.4	16.3 14.6			
Men	32.7	28.4			
Education		Coun	e Group		
		Low	Middle	High	
All Ages	5	28.5	22.2	18.9	
Ages 25-	+ (Non-missing Educ.)	33.0	25.7	20.7	
Ages 25-	+				
Less that	n Secondary	31.8	19.8	12.2	
Seconda	ary Completed 37.3 29.3 23.4			23.4	
More that	in Secondary 39.5 31.7 26.9				

#### **Employent versus workweek**

(a) Employment Rate



1.29

#### Within country patterns

(a) Men

N. 5 ·USA USA-1991 · GBR 0 LTEAUN CZEN CALL CHE •BGR β. - MWI ·ROM ·SVK ·PRISA-FINS • RWA • KHM USA-1895 ·ESP -.PER .BWAPOL • VNM • TZA .COL . TLS PAK · GRCP · MUS · IDN · KEGHA · MNGECU N · UGA ·IRQ °.' •TUR 11 7 7.5 9.5 10 10.5 8 8.5 9 In(GDP per Capita) (b) Women N. •GBR USA-1991 BEL .USA 5 ·NLD ·AUT USA-1973 0 ·BGR ·LVABS/K · PAK •PRT • RWA β ·ROM .LTU ·CYP • MWI -.BWAMUS · VNMMM .TLS IDN BER • UGA USA-1895 •KEN •TZA •GHA • ITA • ECU . POL 2. · GRC • MNG .TUR 3 ·IRQ 7 7.5 9.5 10 10.5 11 8 8.5 9 In(GDP per Capita)

# **Bick et al. (2019)**

#### WHY ARE AVERAGE HOURS WORKED LOWER IN RICHER COUNTRIES?

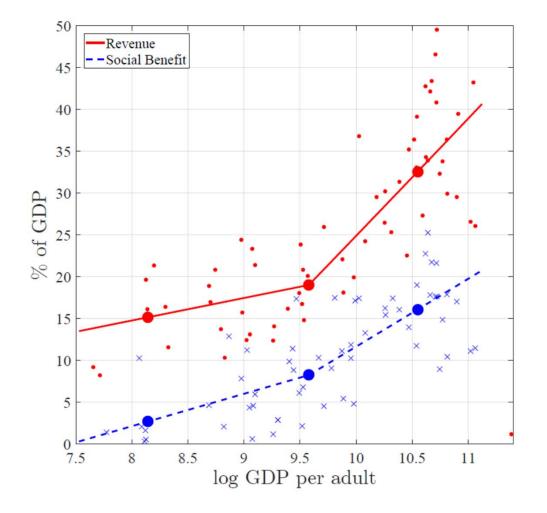
Alexander Bick Nicola Fuchs-Schündeln David Lagakos Hitoshi Tsujiyama

Working Paper 26554 http://www.nber.org/papers/w26554

NATIONAL BUREAU OF ECONOMIC RESEARCH 1050 Massachusetts Avenue Cambridge, MA 02138 December 2019

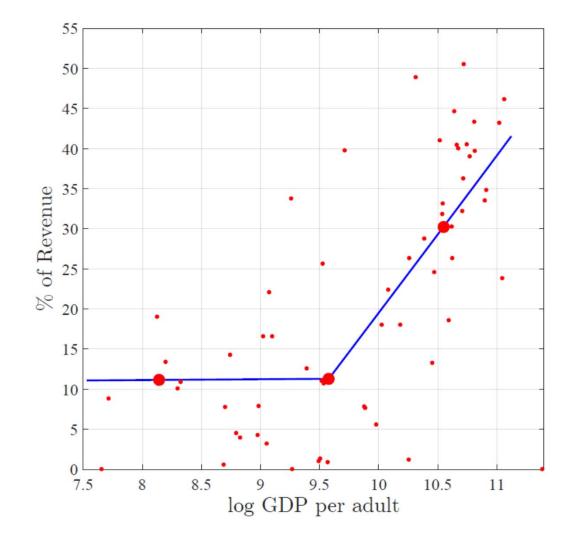
# Bick et al. (2019), cont. (Wagner's Law)

(c) Government Revenue and Social Benefits Relative to GDP



### Bick et al. (2019), continued again

(b) Share of Government Revenues Coming from Labor Income Taxation



# **Chang-Kim, with two-earner family**

- Allow for family
- Income processes assumed orthogonal

$$U = \max_{\{c_t, h_{mt}, h_{ft}\}_{t=0}^{\infty}} E_0 \left\{ \sum_{t=0}^{\infty} \beta^t u(c_t, h_{mt}, h_{ft}) \right\}$$

with

(1) 
$$u(c_t, h_{mt}, h_{ft}) = 2\ln(0.5c_t) - B_m \frac{h_{mt}^{1+1/\gamma}}{1+1/\gamma} - B_f \frac{h_{ft}^{1+1/\gamma}}{1+1/\gamma}$$

(2)

$$V_{ee}(a, x_m, x_f; \lambda, \mu) = \max_{a' \in \mathcal{A}} \left\{ u(c, \bar{h}, \bar{h}) + \beta E \Big[ \max \left\{ V_{ee}', V_{en}', V_{ne}', V_{nn}' \right\} \, \big| \, x_m, x_f, \lambda \Big] \right\}$$

subject to

$$c = w(x_m\bar{h} + x_f\bar{h}) + (1+r)a - a'$$
$$a' \ge \bar{a}$$

1.34

# **Chang & Kim parameters**

#### TABLE 2

#### PARAMETERS OF THE BENCHMARK ECONOMY

Parameter	Description		
$\alpha = 0.64$	Labor share in production function		
$\beta = 0.9807392$	Discount factor		
$\gamma = 0.4$	Intertemporal substitution elasticity		
$B_m = 93.5$	Utility parameter for male		
$B_f = 150.1$	Utility parameter for female		
$\bar{h} = 1/3$	Amount of labor supply when working		
$\rho_x = 0.948 \ (0.925)$	Persistence of productivity <i>x</i> for male (female)		
$\sigma_x = 0.269 \ (0.319)$	Standard deviation of $\epsilon_x$ for male (female)		
$\bar{a} = -4.0$	Borrowing constraint		

#### Some steady-state features

#### CPS Model I Model II **Employment** rates Male 77.33 77.34 77.36 Female 49.75 49.78 49.75 63.54 63.56 63.56 Aggregate Fraction of households 43.87 45.83 Both members working 45.35 Only male working 33.46 31.98 31.52 Only female working 5.88 4.42 3.92 Neither working 16.79 18.23 18.72

LABOR-MARKET STEADY STATES

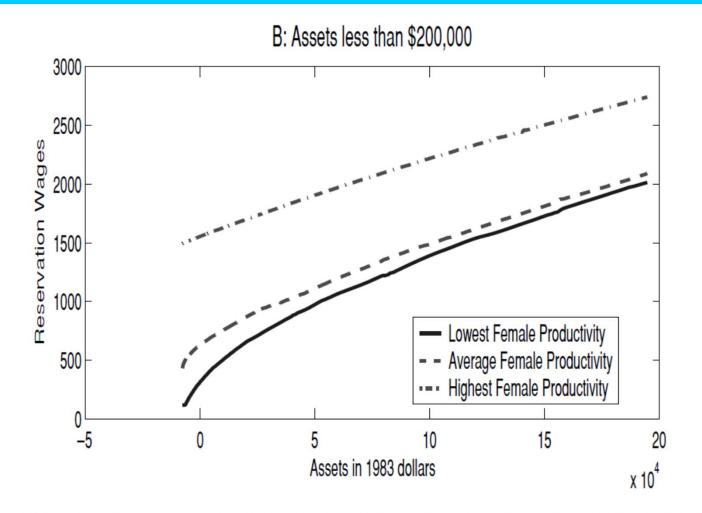
NOTE: All variables are percentages. The statistics for the CPS are annual averages of married households in the March Supplements for the period of 1968–2001.

	PSID	Model I	Model II
Wealth	0.76	0.64	0.61
Earnings	0.53	0.57	0.54

#### GINI INDICES FOR WEALTH AND EARNINGS

NOTE: The PSID statistics reflect the family wealth and earnings in the 1984 survey.

#### **Reservation wages for men**

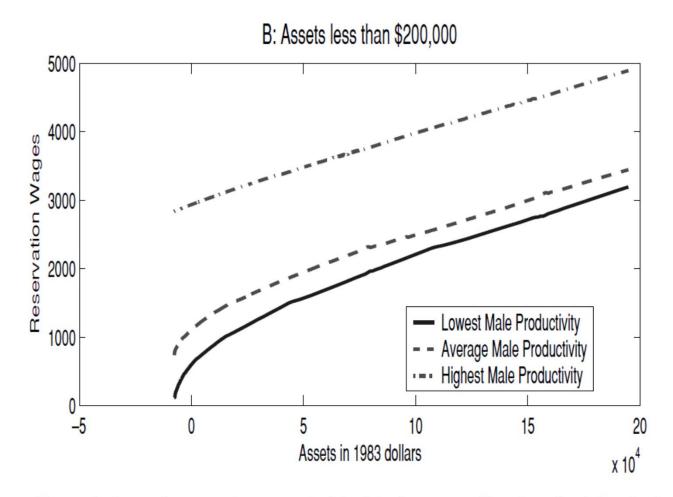


Notes: The graphs denote the reservation-wage schedule of the three types of male worker (whose wife has the highest, average, and lowest productivity). Wages (quarterly earnings) and assets are in 1983 dollars.

#### FIGURE 3

#### RESERVATION-WAGE SCHEDULE OF MALE WORKERS: MODEL I

### **Reservation wages for women**

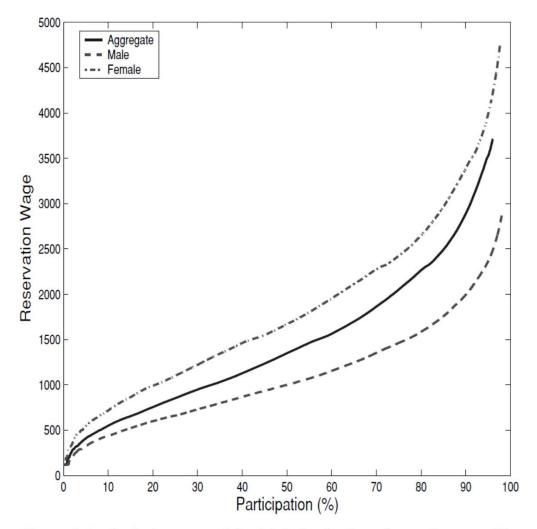


Notes: The graphs denote the reservation-wage schedule of the three types of female worker (whose husband has the highest, average, and lowest productivity). Wages (quarterly earnings) and assets are in 1983 dollars.

#### FIGURE 4

#### RESERVATION-WAGE SCHEDULE OF FEMALE WORKERS: MODEL I

### **Reservation wage schedules**



Notes: The graph denotes the inverse cumulative distribution functions of reservation wages. Wages are quarterly earnings in 1983 dollars.

#### FIGURE 5

#### RESERVATION WAGES AND PARTICIPATION RATES: MODEL I

### **Implied Frisch Elasticities at Extensive Margin**

#### IMPLIED ELASTICITY FROM THE STEADY-STATE RESERVATION-WAGE DISTRIBUTION

Model	Male	Female	Aggregate
Model I	0.84	1.36	0.94
Model II	0.96	1.71	1.12

NOTE: The numbers reflect the elasticity of the labor-market participation rate with respect to reservation wage (evaluated around the steady state) based on the steady-state reservationwage distribution.

### Adjusting wedge for heterogeneous workers

Assume:

- 75% of movements in total hours are via employment (data)
- "Marginal" workers less productive by one third (Barsky, Parker, Solon)
  - biases labor productivity countercyc: add back (3/4)\*(1/3) = 1/4
- Leave workforce causes drop of one-sixth (16.7%) in consumption
  - biases consumption procyc: subtract back -(3/4)\*(1/6) = 1/8

### Labor wedge "corrected" for heterog.

	Elasticity with respect to Total Hours:					
	Uncorrected	Corrected				
Labor Productivity	- 0.33 (.08)	- 0.08				
Total hours	1	1				
Consumption	0.43 (.06)	0.30				
Wedge	-2.19 (.11)	- 1.68				

Notes: Total hours and labor productivity is for private economy; Consumption is nondurables and services. Sample covers 1987 to 2018. All series are logged and HP-filtered. The wedge assumes an IES of 0.5 and a Frisch of 1.0. Correction assumes: (i) three-quarters of movements in total hours via employment; (ii) workers coming in and out of workforce cyclically are one-third less productive; (iii) consumption rises (fall) by one-sixth when enter (exit) workforce.

### Park: "Consumption, Reservations Wages, and Aggregate Labor Supply

Uses empirical joint distribution of wages and consumption to estimate supply elasiticity

Key insight—consumption is sufficient statistic for wealth and future earnings

### **Reservation wage curve**

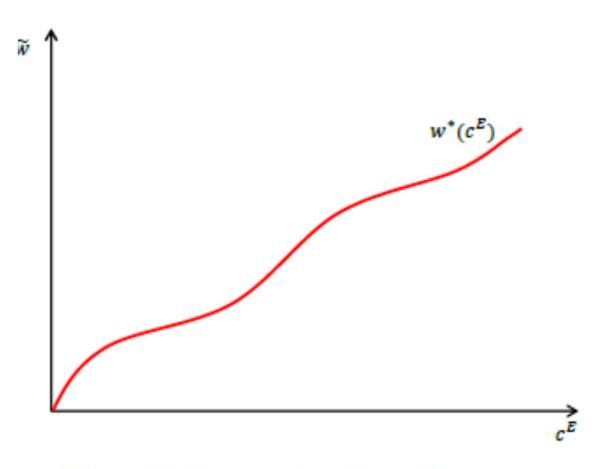


Figure 2: Reservation Wage Curve

### **Robustness of Conditioning on Consumption**

The key result is that the reservation wage conditional on consumption is unique and independent of the state variables. Intuitively, individual saving decisions reflect their future expectations based on constraints they face, and these are all summarized in their consumption choices. Thus, given consumption and wage, the period utility cost of working determines whether they work or not.<sup>9</sup>

Within the class of models defined in equation (1), the reservation property conditional on consumption is robust to the following specification choices:

- 1. arbitrary heterogeneity in discount factors, borrowing constraints, and wage processes,
- 2. time horizon: infinite time versus life cycles,
- 3. flexibility of hours choices: both margins of labor supply versus indivisible labor,
- 4. separability between consumption and leisure in the period utility function,
- 5. two earner's problem when labor is indivisible.

### **Distributions**

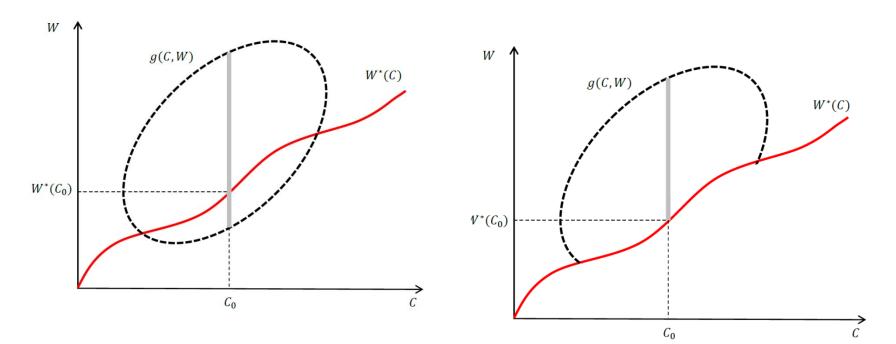


Figure 3: Population Joint Distribution and Employment Rate

Figure 4: Observed Distribution of Workers

# **Observed wages at a particular consumption and the extensive Frisch**

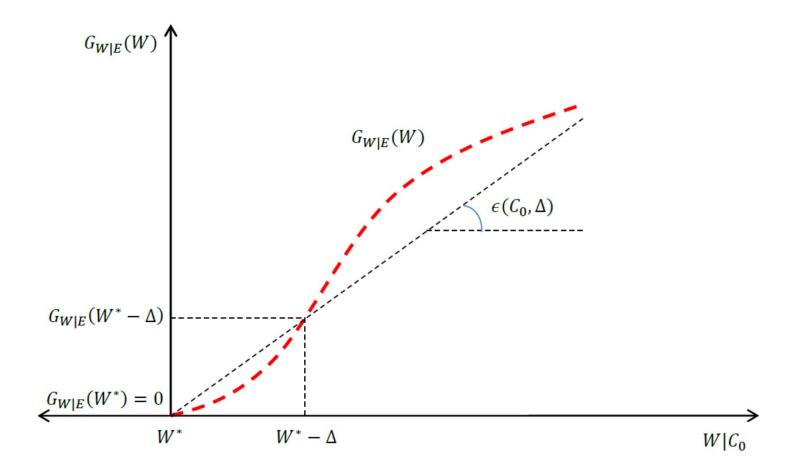


Figure 5: Conditional Distribution and Extensive Margin Frisch Elasticity

### **Allowing for measurement error**

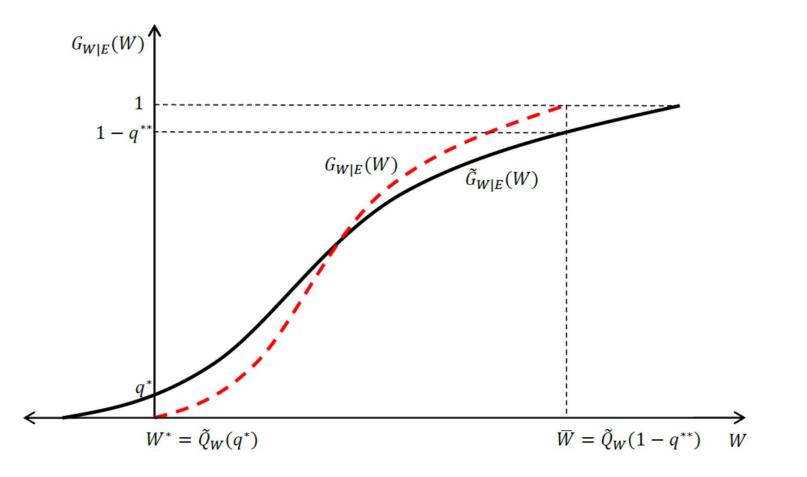


Figure 7: The Effects of Measurement Error on the Distribution

#### **Results for extensive Frisch**

				$q^* = q^{**}$		
		0.03	0.04	0.05	0.06	0.07
Aggr	egate	0.22	0.31	0.40	0.48	0.54
inggrogate		(0.02)	(0.03)	(0.04)	(0.05)	(0.05)
Gender	Male	0.21	0.28	0.34	0.40	0.45
		(0.01)	(0.01)	(0.02)	(0.02)	(0.02)
	Female	0.23	0.34	0.52	0.57	0.64
		(0.03)	(0.05)	(0.07)	(0.08)	(0.08)
		(0.00)	(0.00)	(0.01)	(0.00)	(0.00)
Age	[25, 34]	0.26	0.35	0.43	0.55	0.65
ngo	[20,04]	(0.03)	(0.04)	(0.05)	(0.06)	(0.08)
		(0.03)	(0.04)	(0.05)	(0.00)	(0.00)
	[35, 54]	0.19	0.29	0.39	0.45	0.50
	[33, 34]	(0.02)	(0.29)	(0.04)	(0.45)	(0.05)
		(0.02)	(0.02)	(0.04)	(0.05)	(0.05)
		0.92	0.91	0.41	0.47	0.49
	[55, 65]	0.23	0.31	0.41	0.47	0.48
		(0.03)	(0.04)	(0.06)	(0.06)	(0.06)
	NT II	0.05	0.95	0 51	0.57	0.04
Education	Non-college	0.25	0.35	0.51	0.57	0.64
		(0.03)	(0.04)	(0.06)	(0.07)	(0.06)
	~ 11	0.10	0.00	0.00		0.10
	College	0.16	0.23	0.28	0.34	0.40
		(0.01)	(0.02)	(0.02)	(0.03)	(0.03)
Race	White	0.20	0.28	0.36	0.44	0.52
		(0.01)	(0.02)	(0.03)	(0.04)	(0.04)
	Non-white	0.36	0.52	0.61	0.65	0.67
		(0.06)	(0.08)	(0.08)	(0.10)	(0.13)
Marital	Married	0.21	0.29	0.36	0.44	0.50
Status		(0.01)	(0.02)	(0.03)	(0.03)	(0.03)
	Single	0.25	0.38	0.59	0.65	0.68
		(0.06)	(0.08)	(0.11)	(0.10)	(0.12)

#### Krusell, Mukoyama, Rogerson and Sahin

- Look at Labor Supply Response to Cyclical Fluctuations
- Reflects both comparative advantage and search frictions
   Unlike most DMP models, allow wealth effect
- Give rich depiction of labor flows
  - Job-to-job, exogenous and endogenous separations from employment
  - Endogenous search: transitions between employed, unemployed, OLF
- Movements between unemployed/OLF give insight into labor supply (i.e. substitution and wealth effects), not imposing competitive labor demand

# **Unemployed vs. OLF**

Employed: In the Current Population Survey (CPS), classified as employed if, during the survey <u>reference week</u>, they meet any of the following criteria:

- worked at least 1 hour as a paid employee
- worked at least 1 hour in their own business, profession, trade, or farm
- were temporarily absent from their job, business, or farm
- worked without pay for min. 15 hours in business/farm owned by member of family

Unemployed: In CPS, classified as unemployed if meet all of the following criteria:

- They were not <u>employed</u> during the survey <u>reference week</u>.
- They were available for work during the survey reference week, except for temporary illness.
- They made at least one specific, active effort to find a job during the 4-week period ending with the survey reference week OR they were temporarily laid off and expect to be recalled.
- Classification as unemployed in no way depends upon a person's eligibility for, or receipt of, unemployment insurance benefits.

# **Search recognized**

Active job search methods are defined as those that have the potential to result in job offer without further action on the part of the job seeker. Examples include:

- contacting an employer directly about a job
- having a job interview
- submitting a resume or application to an employer or to a job website
- using a public or private employment agency, job service, placement firm
- contacting a job recruiter or head hunter
- seeking assistance from friends, relatives, or via social networks
- placing or answering a job advertisement
- checking union or professional registers
- Methods that do not constitute an active job search are referred to as passive job search methods. Passive methods are those that could not result in a job offer unless additional steps were taken. Examples include simply looking at job postings or taking a training course.

# **Data: Stocks**

Dataset: Current Population Survey 1978Q1-2012Q3.

	u	lfpr	E
std(x)	0.1170	0.0026	0.0099
corrcoef(x, Y)	-0.84	0.21	0.83
$corrcoef(x, x_{-1})$	0.93	0.69	0.92

- Unemployment rate is countercyclical.
- Labor force participation rate is weakly procyclical.
- Employment rate is procyclical.

#### **Data: Averages of gross worker flows**

Unadjusted Data			Abowd-Zellner Correction				
FROM		ТО		FROM		ТО	
	E	U	N		E	U	N
E	0.957	0.015	0.028	E	0.972	0.014	0.014
U	0.254	0.535	0.211	U	0.228	0.637	0.135
N	0.047	0.028	0.925	N	0.022	0.021	0.957

 Correcting for misreporting reduces flows between U and N, but there still are large flows.

# **Abowd-Zellner (1985) correction**

$$\begin{bmatrix} \hat{E} \\ \hat{U} \\ \hat{N} \end{bmatrix}_{t} = \underbrace{\begin{bmatrix} 1 - \varepsilon_{EU} - \varepsilon_{EN} & \varepsilon_{UE} & \varepsilon_{NE} \\ \varepsilon_{EU} & 1 - \varepsilon_{UE} - \varepsilon_{UN} & \varepsilon_{NU} \\ \varepsilon_{EN} & \varepsilon_{UN} & 1 - \varepsilon_{NE} - \varepsilon_{NE} \end{bmatrix}}_{\mathbf{E}} \begin{bmatrix} E \\ U \\ N \end{bmatrix}_{t}$$

$$\widehat{VN}_t \approx (1 - \varepsilon_{UN} - \varepsilon_{NU}) UN_t + \varepsilon_{UN} UU_t + \varepsilon_{NU} NN_t, \text{ and}$$
$$\widehat{NU}_t \approx (1 - \varepsilon_{UN} - \varepsilon_{NU}) NU_t + \varepsilon_{UN} UU_t + \varepsilon_{NU} NN_t.$$

### **Abowd-Zellner cont.**

$$\mathbf{N}_{t} = \begin{bmatrix} EE & UE & NE \\ EU & UU & NU \\ EN & UN & NN \end{bmatrix}_{t}$$

$$\mathbf{N}_{t} = \mathbf{E}^{-1} \widehat{\mathbf{N}}_{t} \left( \mathbf{E}^{-1} \right)'$$

### **Abowd-Zellner estimates**

Table 1: Abowd and Zellner (1985) estimates of classification errors

Original	Status determined on reinterview							
interview status	Employed	Unemployed	Non-participant					
Employed	98.78	1.91	0.50					
Unemployed	0.18	88.57	0.29					
Non-participant	1.03	9.52	99.21					

Source: Abowd and Zellner (1985, Table 6).

# **Cyclicality of Flows**

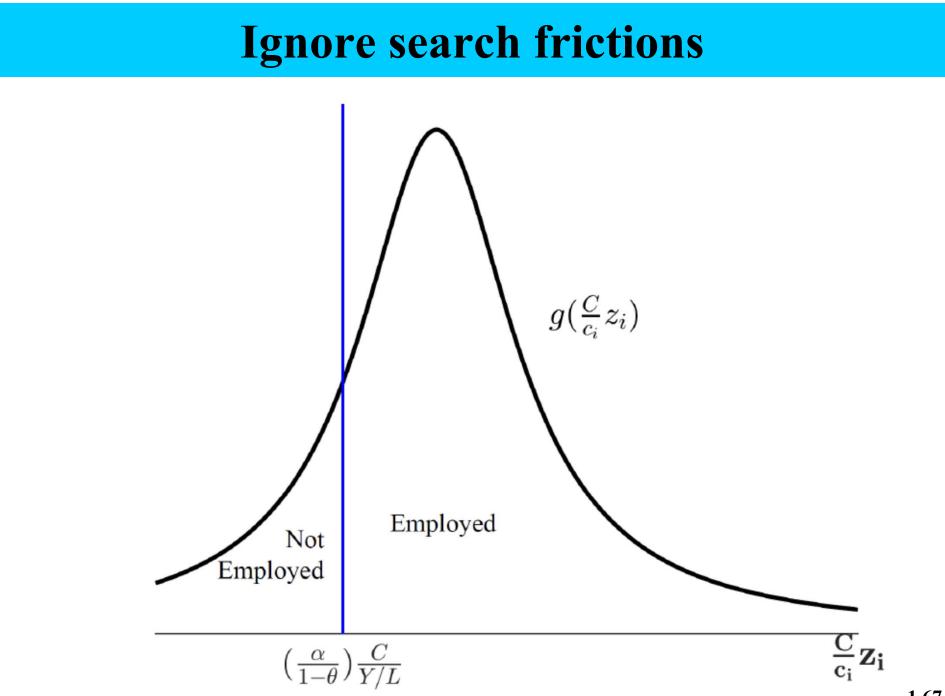
Handlington Data

Unadjusted Data								
	$f_{EU}$	$f_{EN}$	$f_{UE}$	$f_{UN}$	$f_{NE}$	$f_{NU}$		
std(x)	0.075	0.033	0.077	0.053	0.041	0.064		
corrcoef(x, Y)	-0.70	0.35	0.79	0.66	0.61	-0.70		
$corrcoef(x, x_{-1})$	0.69	0.22	0.82	0.71	0.52	0.78		
	Abowc	l-Zellner	Correct	ion				
	$f_{EU}$	$f_{EN}$	$f_{UE}$	$f_{UN}$	$f_{NE}$	$f_{NU}$		
std(x)	0.089	0.083	0.088	0.106	0.103	0.072		
corrcoef(x, Y)	-0.63	0.43	0.76	0.61	0.52	-0.23		
$corrcoef(x, x_{-1})$	0.59	0.29	0.75	0.62	0.38	0.30		

- EU and UE are intuitive, given that the labor demand (frictions from the worker's perspective) is cyclical.
- NE and NU can also be interpreted through the lens of labor market frictions.
- ► *EN* and *UN* are the least intuitive.

# **Model Overview**

- Discrete-time, partial equilibrium model of consumer behavior
- Consumers make consumption/saving decision and labor supply decision.
- Frictional labor market: job offers come with some probability. There are exogenous separations.
- No insurance markets but can save (self-insure).
- Incorporates on-the-job search and realistic unemployment insurance system.

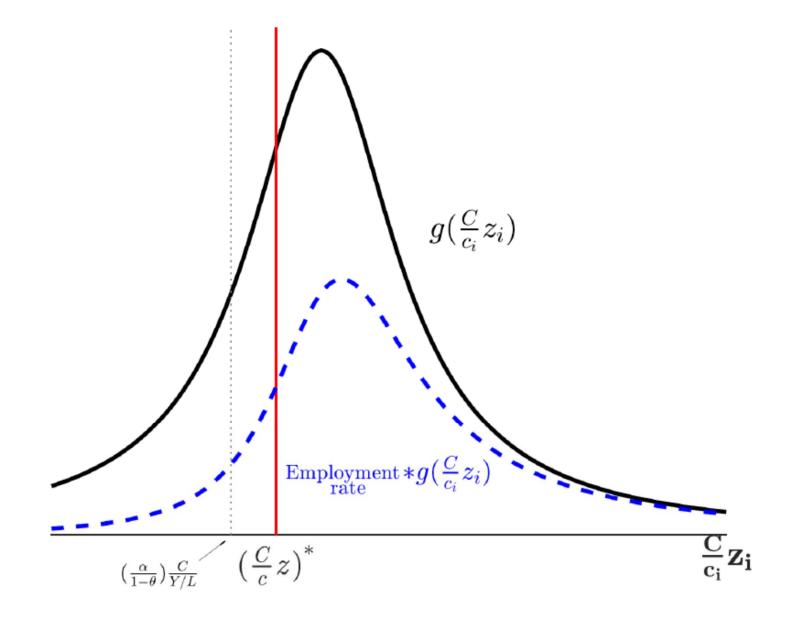


<sup>1.67</sup> 

#### Add rich set of frictions, choices governing flows

- Two exogenous flows: Exogenous job-to-job and exogenous separations
- Endogenous separations, endogenous take up of offers, endogenous search
- A number of qualitative predictions:
  - Separation rate decreasing in  $\frac{w_i}{c_i}$
  - Probability unemployed versus OLF increasing in  $\frac{w_i}{c_i}$
  - Probability transit unemployed to employed increasing in  $\frac{w_i}{c_i}$
- $\frac{C}{Y/L}$  remains key statistic for judging cyclicality of endogenous choices

#### **Comparative advantage Gets "smoothed" out**



#### **Comp. adv. still key to search/separation decisions**

- Focus on search decision
  - Take perfect-insurance case: critical  $z^*$ , search if, only if,  $z \ge z^*$
  - perturb  $z^*$  today and next period to hold matches constant going forward

Net benefit (ignoring heterogeneity in match quality)

$$= -\gamma + (\lambda_u - \lambda_n) \left( \frac{z_i(1-\theta)Y}{L} \frac{1}{C} - \alpha + (1-\sigma)\beta \frac{\gamma}{(\lambda_u - \lambda_n)} \right)$$
$$\Rightarrow z^* = \left( \frac{\alpha}{1-\theta} \right) \frac{C}{Y/L} \left( 1 + \frac{\gamma/\alpha}{(\lambda_u - \lambda_n)} \left( 1 - \beta(1-\sigma) \right) \right)$$

- Threshold (as without frictions) dictated by statistic:  $\frac{Y/L}{C} = \frac{(Y/E)(E/L)}{C}$ 

#### Steady state calibration

- 1 period = 1 month
- Set  $\beta$  and  $\tau$ :  $\beta = 0.9947$  and  $\tau = 0.3$ .
- ▶  $\rho = 0.996$  and  $\sigma_{\varepsilon} = 0.096$ , from the micro estimates of earnings processes.
- ► r, w, and T come from the background general equilibrium model.
- ► UI parameters: µ = 1/6 (eligible for 6 months) and replacement rate= 0.23, to match the UI/(total earnings).
- $\bar{\gamma} = (3.5/40) \times \alpha$ , from time use data.

#### Steady state calibration

- 1 period = 1 month
- Set  $\beta$  and  $\tau$ :  $\beta = 0.9947$  and  $\tau = 0.3$ .
- ▶  $\rho = 0.996$  and  $\sigma_{\varepsilon} = 0.096$ , from the micro estimates of earnings processes.
- ► r, w, and T come from the background general equilibrium model.
- ► UI parameters: µ = 1/6 (eligible for 6 months) and replacement rate= 0.23, to match the UI/(total earnings).
- $\bar{\gamma} = (3.5/40) \times \alpha$ , from time use data.

For the following, we use the steady state version of the model to find the (average value of) parameter values.

- $\alpha$  to match E.
- $\lambda_u$  to match the unemployment rate.
- $\lambda_n$  to match the NE flow rate.
- $\sigma$  to match the EU flow rate.
- λ<sub>e</sub> and σ<sub>q</sub> to match job-to-job transition rate and the wage gain upon transition.

Steady state flows (monthly)

AZ-Adjusted Data				Mo	del			
FROM		ΤO		FROM		ΤO		
	E	U	N		E	U	N	
E	0.972	0.014	0.014	E	0.972	0.014	0.014	
U	0.228	0.637	0.135	U	0.219	0.652	0.130	
N	0.022	0.021	0.957	N	0.022	0.020	0.958	

Gross Worker Flows in the Data and the Model

Each flow is matched well to the data.

 The model captures the relationship between wealth and flows well.

wealth and flows

#### Business cycle analysis

- We let the labor market condition (λ<sub>u</sub>, λ<sub>n</sub>, λ<sub>e</sub>, σ) vary over the business cycle. We assume a two-point Markov process (all move together) that resembles business cycles.
- $\lambda_u$  fluctuates so that the standard deviation of  $f_{UE}$  to the data.
- $\sigma$  fluctuates so that the standard deviation of  $f_{EU}$  to the data.
- $\lambda_n$  and  $\lambda_e$  maintain the same proportion to  $\lambda_u$ .
- w and r are constant.

#### Cyclical properties of stocks

	Data			Model			
	u	lfpr	E	u	lfpr	E	
std(x)	0.1170	0.0026	0.0099	0.1207	0.0015	0.0096	
corrcoef(x, Y)	-0.84	0.21	0.83	-0.99	0.37	0.995	
$corrcoef(x, x_{-1})$	0.93	0.69	0.92	0.87	0.71	0.89	

- The labor force participation rate is weakly procyclical. Two forces:
  - In recession, when the job-finding rate decreases and the separation rate increases, workers increases participation rate to offset the frictions (the wealth effect of labor supply). This is a countercyclical force.
  - The return to search is larger in booms, because (i) λ<sub>u</sub> is larger, (ii) σ is smaller, and (iii) λ<sub>e</sub> is larger (implicit 'wage movement'). This is a procyclical force.
  - These two forces almost offset with each other.

#### Cyclical properties of gross flows

Data							
	$f_{EU}$	$f_{EN}$	$f_{UE}$	$f_{UN}$	$f_{NE}$	$f_{NU}$	
std(x)	0.089	0.083	0.088	0.106	0.103	0.072	
corrcoef(x, Y)	-0.63	0.43	0.76	0.61	0.52	-0.23	
$corrcoef(x, x_{-1})$	0.59	0.29	0.75	0.62	0.38	0.30	
		Mod	el				
	$f_{EU}$	$f_{EN}$	$f_{UE}$	$f_{UN}$	$f_{NE}$	$f_{NU}$	
std(x)	0.089	0.057	0.088	0.029	0.051	0.076	
corr(x, Y)	-0.79	0.21	0.69	0.47	0.57	-0.96	
$corr(x, x_{-1})$	0.76	0.21	0.70	0.34	0.66	0.87	

- EU and UE flows come from  $\lambda_u$  and  $\sigma$  fluctuations.
- NE and NU are affected by  $\lambda_n$  fluctuations.

#### Cyclical properties of gross flows

- ▶ Procyclical *EN*: Within *E*, there are more workers at the boundary of participation in booms. There are more 'recently employed' workers who are close to the boundary, because in booms there are more movement from *U* and *N* to *E*.
- Procyclical UN: Within U, there are more workers at the boundary of participation in booms. In recession, there are more 'poor' and 'high-productivity' workers who have been in U for a long time, and has a large z.
- The labor supply intuition (except for the wealth effect) goes the opposite direction, as working becomes more attractive in booms.

### **The Labor Wedge and Labor Demand**

# Bils, Klenow, Malin: "Resurrecting ..."

#### Decomposing the Labor Wedge

Hours worked appear to be inefficiently low in recessions.

• Labor Wedge is high:  $\mu \equiv \frac{mpn}{mrs}$ 

Labor Wedge is the product of:

**1** Labor Market Wedge: 
$$\mu^w \equiv \frac{w/p}{mrs}$$

**2** Product Market Wedge: 
$$\mu^p \equiv \frac{mpn}{w/p} \equiv \frac{p}{mc}$$

Uses (aggregate) wage data

- E.g., Gali, Gertler, Lopez-Salido (2007), Karabarbounis (2014)
- Measure of Price of Labor: w/p = average wage
- Key Assumption: all workers employed in spot markets.
- Conclusion:  $\mu^{w}$  accounts for nearly all cyclicality of  $\mu$ .

BUT, conclusion depends critically on wage measure used.

- Alternative theories emphasize durable nature of employment and wage smoothing.
- w/p can be much more procyclical using other wage measures.

Decomposes Labor Wedge  $\mu$  without using wage data.

Recall:  $\mu^{p} \equiv \frac{p}{mc}$ 

Consider 2 alternative inputs:

# 1 Self-Employed • $\frac{p}{mc} = \frac{p}{p \cdot mrs/mpn} = \frac{mpn}{mrs}$ , or $\mu^p = \mu$

2 Intermediate Inputs

$$\frac{p}{mc} = \frac{p}{p_m/mpm}$$

Our point estimates:  $\mu^{p}$  accounts for the cyclical variation in  $\mu$ 

- Self-Employed  $\mu$  is just as cyclical as all-worker  $\mu$
- Intermediate Inputs  $\mu^{p}$  is just as cyclical as  $\mu$

Thus, countercyclical price markups deserve a central place in business cycle research, alongside labor market frictions.

Measuring the Labor Wedge

- Focus on Intensive Margin
- Decompose using Wage Data

Our 2 Alternative Decompositions

1 Self-Employed

**2** Intermediate Inputs

#### **Representative-Agent Labor Wedge**

#### Preferences:

$$\mathbb{E}_{0} \sum_{t=0}^{\infty} \beta^{t} \left\{ \frac{c_{t}^{1-1/\sigma}}{1-1/\sigma} - \nu \frac{n_{t}^{1+1/\eta}}{1+1/\eta} \right\}$$

Production:

$$y_t = z_t k_t^{\alpha} n_t^{1-\alpha}$$

Labor Wedge:

$$\begin{aligned} \ln(\mu_t) &\equiv \ln(mpn_t) - \ln(mrs_t) \\ &= \ln\left(\frac{y_t}{n_t}\right) - \left[\frac{1}{\sigma}\ln(c_t) + \frac{1}{\eta}\ln(n_t)\right] \end{aligned}$$

#### Extensive and Intensive Margin Labor Wedges

- Consider extensive and intensive margins of labor supply
- Why?
  - Can base Frisch elasticity on micro estimates using hours margin
  - Self-employed wedge will be on intensive margin only
  - Product market distortions should impact wedge on both margins
    - If wedge is only important on one margin, product market distortions must have little cyclical importance.

### Theory with Both Extensive and Intensive Margins

Preferences:

$$\mathbb{E}_0 \sum_{t=0}^{\infty} \beta^t \left\{ \frac{c_t^{1-1/\sigma}}{1-1/\sigma} - \nu \left( \frac{h_t^{1+1/\eta}}{1+1/\eta} + \psi \right) \boldsymbol{e}_t \right\}$$

Production:

$$y_t = z_t k_t^{\alpha} (e_t h_t)^{1-\alpha}$$

Search Frictions:

- Matching Technology:  $m_t = v_t^{\phi} f(u_t)$
- Vacancy-posting cost: κ
- Separation rate:  $\delta$

$$\begin{aligned} &ln(\mu_t) &\equiv ln(mpn_t) - ln(mrs_t) \\ &= ln\left(\frac{y_t}{n_t}\right) - \left[\frac{1}{\sigma}ln(c_t) + \frac{1}{\eta}ln(h_t)\right] \end{aligned}$$

- $h_t$  = hours per *worker*
- $\eta = 0.5$
- *σ* = 0.5

#### Cyclicality of Intensive-Margin Labor Wedge

$$ln(\mu_t) = \alpha + \beta \cdot ln(cyc_t) + \epsilon_t$$

Elasticity wrt GDP

Labor Wedge -1.91 (0.13)

Labor Productivity-0.10 (0.08)

Cons per capita 0.61 (0.03)

Hours per worker 0.30(0.07)

• Quarterly data, 1987-2012 with  $\sigma = 0.5, \eta = 0.5$ 

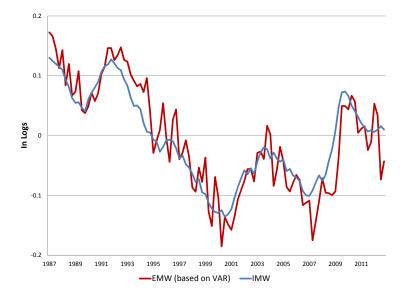
Consider spending today to generate one more matched worker, then reduce spending next period to cut matches by  $1 - \delta$  workers:

*EMW*  $\approx \ln(y/n) - 1/\sigma \cdot \ln(c) - dynamic cost of vacancy matching$ 

So:

$$EMW - IMW = \frac{1}{\eta} In(h) - dynamic \ cost \ of \ vacancy \ matching$$

## EMW vs. IMW



	Elasticity wrt		
	GDP	Total Hours	
EMW	-1.89 (0.28)	-1.54 (0.15)	
IMW	-1.91 (0.13)	-1.38 (0.05)	

- Quarterly data, 1987-2012
- *σ* = 0.5, *η* = 0.5
- $\delta = 0.105, r = 0.004, \phi = 0.5, \frac{\kappa v}{m} = 0.4, \gamma = 0.16$
- · Expectational terms in EMW constructed using VAR approach

#### Decomposing the Wedge

Decomposition:

$$\ln(\mu_t) = \left[ ln\left(\frac{y_t}{n_t}\right) - ln\left(\frac{w_t}{p_t}\right) \right] + \left[ ln\left(\frac{w_t}{p_t}\right) - \frac{1}{\sigma}ln(c_t) - \frac{1}{\eta}ln(h_t) \right]$$
$$= ln(\mu_t^p) + ln(\mu_t^w)$$

Cyclicality:

$$\begin{aligned} & \ln(\mu_t) = \alpha + \beta \cdot \ln(cyc_t) + \epsilon_t \\ & \ln(\mu_t^p) = \alpha^p + \beta^p \cdot \ln(cyc_t) + \epsilon_t \\ & \ln(\mu_t^w) = \alpha^w + \beta^w \cdot \ln(cyc_t) + \epsilon_t \end{aligned}$$

Note:  $\beta = \beta^{p} + \beta^{w}$ .

## Wedge Decomposition: Standard Approach

Elasticity wrt GDP

$$\mu$$
 -1.91 (0.13)

$$\mu^{p}\left(\frac{w}{p} = AHE\right) \qquad -0.04 \ (0.13)$$

#### Semi-elasticities wrt the Unemployment Rate (s.e.'s):

Average Hourly Earnings	-1.8 (0.7)
New-hire Wage	-3.0 (0.8)
User Cost of Labor	-5.2 (0.8)

Source: Kudlyak (2015) using the NLSY

# Kudlyak (2014) User Cost of Labor

- Question? What is wage cost of employing one more worker today?
- Typically treated as the flow wage rate, typically measured by average hourly earning
- But employment relationships are often long-lasting: wages reflect installment payments (Lazear, 1976, Hall, 1980)
- User cost is impact of hiring now, vs. next period on PDV of wage bill:

$$UC_t^W \equiv E_t(PDV_t - \beta(1 - \delta)PDV_{t+1}),$$

. . .

$$UC_t^W = w_{t,t} + E_t \sum_{\tau = t+1}^{\infty} (\beta(1-\delta))^{\tau-t} (w_{t,\tau} - w_{t+1,\tau}).$$

# **Measurement based on NLS fixed-effect regressions**

			With industry controls	
	Full sample	New hires only	Full sample	New hires only
	(1)	(2)	(3)	(4)
$u_{current}$	-1.78**	-3.00***	-2.02**	-2.99***
	(0.72)	(0.78)	(0.93)	(0.92)
Grade	$7.98^{***}$	$12.52^{***}$	$7.42^{***}$	$11.67^{***}$
	(1.52)	(1.65)	(1.55)	(1.98)
Experience	$4.22^{**}$	8.28***	$3.71^{*}$	$7.77^{***}$
	(1.66)	(1.75)	(1.84)	(2.10)
$Experience^2$	-0.13***	-0.14***	-0.14***	-0.15***
	(0.02)	(0.02)	(0.02)	(0.02)
Tenure	$3.55^{***}$	4.02	$3.71^{***}$	7.57
	(0.23)	(4.60)	(0.29)	(4.95)
$Tenure^2$	-0.11***	3.29	-0.13***	-0.29
	(0.02)	(4.30)	(0.02)	(4.64)
Trend	1.03	-3.52**	1.55	-2.95
	(1.74)	(1.70)	(1.86)	(2.04)
Union dummy	$0.19^{***}$	$0.17^{***}$	$0.16^{***}$	$0.15^{***}$
	(0.01)	(0.02)	(0.01)	(0.02)
Industry dummies	x	x	yes	yes
-			-	-
Observations	52593	19406	46753	16963
R-squared	0.529	0.472	0.558	0.507

Note – The data in Columns 1 and 2 are from NLSY79, men only, 1978 - 2004. The sample of new hires is restricted to observations with tenure less than 1 year. The dependent variable is the natural logarithm of real hourly wage. All regressions are estimated with fixed effects using sampling weights. Unemployment rate is the annual unemployment rate. Columns 3 and 4 include controls for 14 industries, and are estimated on the 1978 - 2002 sample because of the change in the industry classification between 2002 and 2004. The estimated standard errors are in parentheses, clustered by time. The coefficients and standard errors are multiplied by 100. P-values: \*\*\*p<0.01, \*\* p<0.05, \* p<0.1.

# **Cyclicality of User Cost vs. Ave. Hrly Earnings**

Cyclicality of the user cost of labor.

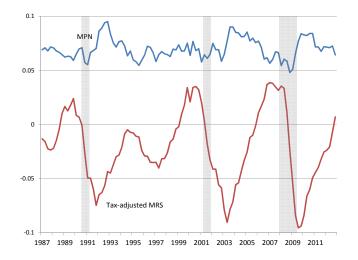
	Semi-elasticity with respect to unemployment		
Measure	Estimate	95% Confidence Interval	
User cost of labor, $UC_t^W$	-5.20*** (0.76)	-6.69 3.71	
Wage of new hires	-3.00*** (0.78)	-4.61 1.40	
Average wage	-1.78** (0.72)	-3.260.30	

*Note*: The semi-elasticity is the coefficient on the unemployment rate from the regression of the (natural logarithm of the) respective series on the contemporaneous unemployment rate and other controls. The estimates for the user cost of labor are from the regression of the (natural logarithm of the) user cost of labor on the unemployment rate and a time trend (annual). There are 20 observations in the regression of the wage component of the user cost – from 1978 to 1997. The bootstrapped standard errors are in parentheses (1000 replications). All coefficients and standard errors are multiplied by 100.

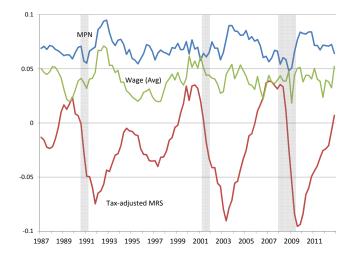
# Robustness

	# of years in calculating $UC_t^W$		
	5  years	7 years	9 years
$UC_t^W,  \delta_t = const$	-5.03	-5.24	-5.33
	(0.77)	(0.81)	(0.83)
$UC_t^W,  \delta_t$	-5.02	-5.19	-5.27
	(0.80)	(0.76)	(0.81)
$UC_t^W,  \delta_{t_0,t}$	-4.79	-4.91	-4.89
	(0.16)	(0.59)	(0.70)

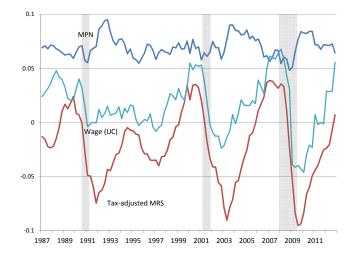
Note - The estimates are from the regression of the natural logarithm of the wage component of the user cost of labor on the unemployment rate and a time trend (annual). There are 18 observations in each regression - from 1978 to 1995. The bootstrapped standard errors are in parentheses (1000 replications). All coefficients and standard errors are multiplied by 100. The three rows reflect different ways of treating the separation rates in the construction of the wage component of the user cost of labor: 1) constant separation rate,  $\delta_t = const$ ; 2) separation rate that depends on the current period,  $\delta_t$ ; and 3) separation rate that depends both on the current period and the period when the job started,  $\delta_{t_0,t}$ .



#### Wedge Decomposition: Avg Wage



#### Wedge Decomposition: User Cost of Labor



## EMW and IMW Decomposition

$$EMW = \left[ ln\left(\frac{y/n}{w/p}\right) - \tilde{S} \right] + \left[ ln\left(\frac{w}{p}\right) + \tilde{S} - S - \frac{1}{\sigma}ln(c) - ln(\Omega) \right],$$
  
where  $\tilde{S} = S$ , but with  $\phi = 1$ .

$$IMW = \left[ ln\left(\frac{y/n}{w/p}\right) \right] + \left[ ln\left(\frac{w}{p}\right) - \frac{1}{\sigma}ln(c) - \frac{1}{\eta}ln(h) \right]$$

Elasticity wrt GDP	EMW	IMW
$\mu$	-1.89 (0.28)	-1.91 (0.13)
$\mu^{p}\left(rac{w}{p}=AHE ight)$	-0.32 (0.13)	-0.04 (0.13)
$\mu^{p}\left(\frac{w}{p}=NH\right)^{\prime}$	-0.98 (0.16)	-0.70 (0.16)
$\mu^{p}\left(\underline{w}_{p}=UC\right)$	-2.17 (0.21)	-1.89 (0.21)

#### Outline

Measuring the Labor Wedge

- Fous on Intensive Margin
- Decompose using Wage Data

Our 2 Alternative Decompositions

1 Self-Employed

**2** Intermediate Inputs

#### Idea:

- Compare the wedge for the self-employed ( $\mu_{se}$ ) to the wedge for all workers ( $\mu$ ).
- Assuming  $\mu_{se} = \mu_{se}^{p} = \mu^{p}$ , comparison yields  $\mu^{p}$  vs.  $\mu$ .

#### Focus on intensive (hours) margin

· Extensive movements could reflect costs of starting business

### Data on Self-Employed

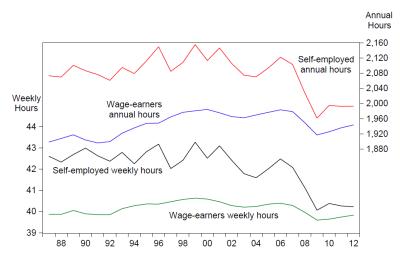
Hours and Earnings: March CPS

- "Self-employed"
  - Primary job is (nonag) self-employment.
  - ▶ 95% of earnings from primary job
- Trim sample to deal with top and bottom coding
- Hours: usual weekly hours (also total annual hours)
- Earnings from primary job
- Examine year-to-year changes for "matched" workers

Consumption: Consumer Expenditure Survey

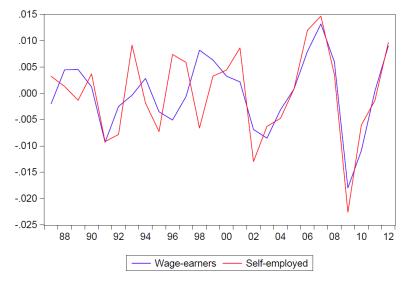
• Construct *relative* consumption of self-employed

### Hours: Self-Emp vs. Wage-Earn (Repeated CPS)



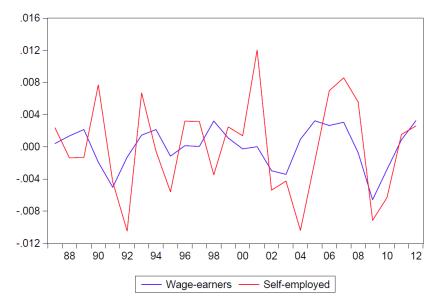
Weekly Hours cyclicality (wrt GDP): 0.37 (0.14), 0.20 (0.02) Annual Hours cyclicality (wrt GDP): 0.57 (0.18), 0.39 (0.04)

#### Annual Hours: Self-Emp vs. Wage Earn (Matched)



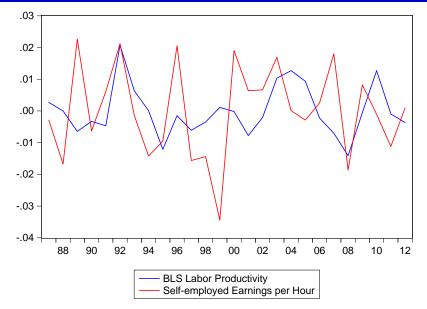
Cyclicality (wrt GDP): 0.54 (0.13), 0.57 (0.07)

### Weekly Hours: Wage-Earn vs. Self-Emp (Matched)



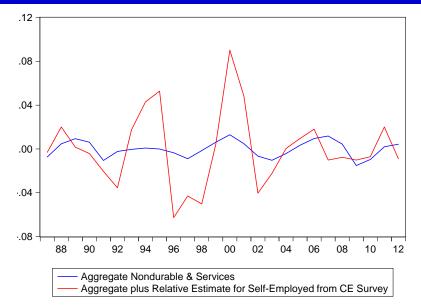
Cyclicality (wrt GDP): 0.17 (0.03), 0.28 (0.07)

#### Productivity: All Workers vs. Self-Emp



Cyclicality (wrt GDP): -0.21 (0.07), -0.13 (0.19)

#### Consumption: All Workers vs. Self-Emp

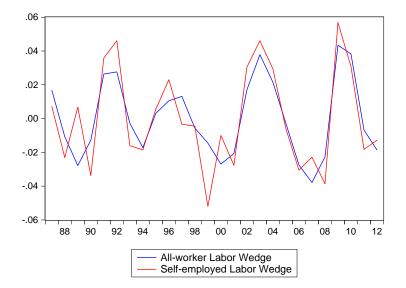


Cyclicality (wrt GDP): 0.64 (0.04), 1.27 (0.56)

## Cyclicality of the Labor Wedge: All vs. Self-Employed

	Labor Wedge			
Elasticity wrt	(1)	(2)	(3)	(4)
Real GDP	-1.87 (0.10)	-2.06 (0.17)	-1.97 (0.25)	-3.23 (1.00)
Hours	All	SE	SE	SE
MPN	Agg. y/n	Agg. y/n	SE earn/hr	SE earn/hr
Consumption	NIPA PCE	NIPA PCE	NIPA PCE	NIPA PCE + CE adj.

#### Labor Wedge for Self-Employed vs. All Workers



(Baseline) self-employed wedge is at least as countercyclical as all-worker wedge.

Robustness:

- 1 Use only *unincorporated* self-employed
- 2 Weight CPS observations by industry
- Weight CPS observations by share of self-employed in industry-occupation that have employees

Conclusion:  $\mu^{p}$  accounts for the bulk of cyclical variation in  $\mu$ .

#### Outline

#### Measuring the Labor Wedge

- Focus on Intensive Margin
- Decompose using Wage Data

Our 2 Alternative Decompositions

- 1 Self-Employed
- **2** Intermediate Inputs

#### Approach 2: Intermediate Inputs

Production function:

$$y = \left[\theta m^{\frac{\varepsilon-1}{\varepsilon}} + (1-\theta) \left[ z_{\nu} \left[ \alpha k^{\frac{\omega-1}{\omega}} + (1-\alpha)(z_n n^{\frac{\omega-1}{\omega}}) \right]^{\frac{\omega}{\omega-1}} \right]^{\frac{\varepsilon}{\varepsilon-1}} \right]^{\frac{\varepsilon}{\varepsilon-1}}$$

c

Marginal Product wrt Intermediates:

$$mpm_t = \theta \left(\frac{y_t}{m_t}\right)^{\frac{1}{\varepsilon}}$$

Product Market Wedge:

$$\mu_t^{p} = \frac{p_t}{mc_t} = \frac{p_t}{p_{mt}/mpm_t}$$

Product Market Wedge

$$\mu_{it}^{p} = \frac{p_{it} y_{it}}{p_{m,it} m_{it}} \left(\frac{y_{it}}{m_{it}}\right)^{\frac{1}{\varepsilon}-1}$$

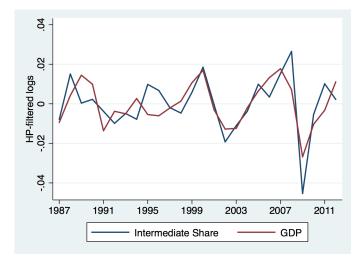
#### BLS Multifactor Productivity Database

- Annual data, 1987-2012
- 60 industries (18 manufacturing)
- Output and KLEMS inputs, nominal and real

Baseline:  $\varepsilon = 1$ 

• Robustness:  $\varepsilon < 1$ 

#### Cyclicality of Intermediate Share



#### Cyclicality of Intermediates-based $\mu^{p}$

$$\ln\left(\mu_{it}^{p}\right) = \alpha_{i} + \beta^{p} \cdot \ln(cyc_{t}) + \epsilon_{it}$$

Elasticity wrt GDPAll Industries-0.94 (0.24)Manufacturing-0.95 (0.32)Non-Manufacturing-0.94 (0.24)

• Baseline estimates with  $\varepsilon = 1$ .

#### Intuition for Intermediates Results

• If *w* and  $p_m$  reflect true shadow prices, then (for  $\varepsilon = 1$ )

$$\frac{w n}{p_m m} = const.$$

 But empirically, intermediate expenditures more procyclical than labor expenditures ⇒ intermediates-based μ<sup>p</sup> is more countercyclical.

$$\ln\left(\mu^{p}\right) = \ln\left(\frac{p y}{p_{m}m}\right) = \ln\left(\frac{p y}{w n}\right) + \ln\left(\frac{w n}{p_{m}m}\right)$$

• Possible reconciliation: w doesn't reflect true shadow price.

## Industry-level Labor Wedge $(\mu_i)$

#### Preferences:

$$\mathbb{E}_{0}\sum_{t=0}^{\infty}\beta^{t}\left\{\frac{\boldsymbol{c}_{t}^{1-1/\sigma}}{1-1/\sigma}-\nu\sum_{i}\left[\left(\frac{\boldsymbol{h}_{it}^{1+1/\eta}}{1+1/\eta}+\psi\right)\boldsymbol{e}_{it}\right]\right\}$$

Marginal Product wrt Labor (for  $\varepsilon = \omega = 1$ ):

$$mpn_{it} = \frac{y_{it}}{n_{it}}$$

Labor Wedge (intensive-margin):

$$\ln(\mu_{it}) = \ln\left(\frac{p_{it} mp_{it}}{p_t mrs_{it}}\right) = \ln\left(\frac{p_{it}}{p_t} \frac{y_{it}}{n_{it}}\right) - \left[\frac{1}{\sigma} \ln(c_t) + \frac{1}{\eta} \ln(h_{it})\right]$$

## Cyclicality of *Industry-level* Labor Wedge $(\mu_i)$

$$ln(\mu_i) = ln\left(\frac{p_i v_i}{p n_i}\right) + ln\left(\frac{y_i}{v_i}\right) - \left[\frac{1}{\sigma}ln(c) + \frac{1}{\eta}ln(h_i)\right]$$

Elasticity wrt GDP

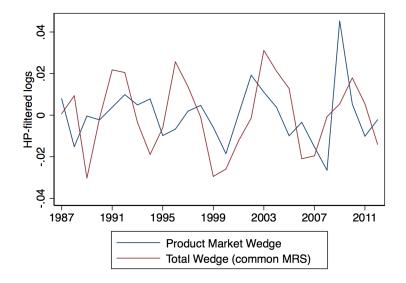
All Industries -0.89 (0.26)

Manufacturing -0.72 (0.39)

Non-Manufacturing -0.93 (0.24)

• Baseline estimates with  $\varepsilon = 1$ .

#### Intermediates-based $\mu^{\rho}$ vs. Total Labor Wedge $\mu$



•  $\varepsilon < 1 \Rightarrow \mu_i^p$  more countercyclical

$$\ln\left(\mu_{it}^{p}\right) = \ln\left(\frac{p_{it} y_{it}}{p_{m,it}m_{it}}\right) + \left(\frac{1}{\varepsilon} - 1\right)\ln\left(\frac{y_{it}}{m_{it}}\right)$$

•  $\varepsilon < 1 \Rightarrow \mu_i$  *less* countercyclical

$$ln(\mu_{it}) = ln\left(\frac{p_{it}}{p_t}\frac{y_{it}}{n_{it}}\right) + \left(\frac{1}{\varepsilon} - 1\right)ln\left(\frac{y_{it}}{v_{it}}\right) - ln\left(mrs_{it}^h\right)$$

•  $\therefore \varepsilon < 1 \Rightarrow \mu^{p}$  accounts for > 100% of cyclicality of  $\mu$ .

#### Measuring the Labor Wedge

- Examine both Extensive and Intensive Margins
- Decompose using Wage Data
- Our 2 Alternative Decompositions
  - 1 Self-Employed
  - **2** Intermediate Inputs

Discuss Other Non-Wage Decompositions

## Other ways to get price markups without wage data

- Capital expenditures (Galeotti and Schiantarelli, 1998)
- Advertising (Hall, 2014)
- Inventories
  - Finished goods inventories
    - Bils and Kahn, 2000
    - Kryvtsov and Midrigan, 2012
  - Work-in-process inventories (appendix)

## Summary of other ways to get price markups

- Capital expenditures  $\Rightarrow$  countercyclical markups
- Advertising  $\Rightarrow$  acyclical markups (maybe)
- Inventories  $\Rightarrow$  countercyclical markups

All involve dynamics, requiring one to measure any adjustment costs and the stochastic discount factor.

Self-Employed and Intermediates require only static measurements.

Our point estimates:  $\mu^{p}$  accounts for the cyclical variation in  $\mu$ 

- Self-Employed  $\mu$  is just as cyclical as all-worker  $\mu$
- Intermediate Inputs  $\mu^{p}$  is just as cyclical as  $\mu$

Countercyclical price markups deserve a central place in business cycle research, alongside labor market frictions.

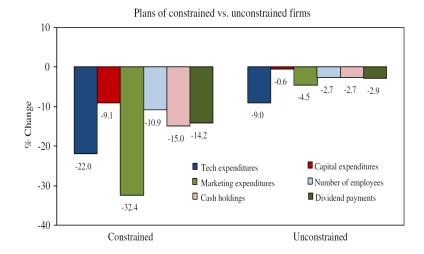
- Sticky prices
- Customer base and/or learning-by-doing + financial shocks
- Countercyclical risk or risk-aversion

# Some approaches to Cyclical L<sup>D</sup>

- Financial Frictions and Fluctuations in Volatility, by Arellano, Bai, and Kehoe
- Asset Prices and Unemployment Fluctuations, by Kehoe, Lopez, Midrigan, and Pastorino

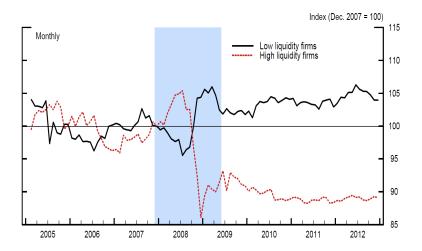
- Cyclicality in Markup, some examples
  - Rotemberg and Woodfood (1999 Handbook)
  - Gilchrist, Schoenle, Sim and Zakrajsek (2015)

#### Evidence for financial constraints



Source: Campello, Graham and Harvey (2010)

#### Evidence on relative prices



Source: Gilchrist, Schoenle, Sim and Zakrajšek (2015)

## Financial Frictions and Fluctuations in Volatility by Arellano, Bai, and Kehoe (2017)

#### Motivation

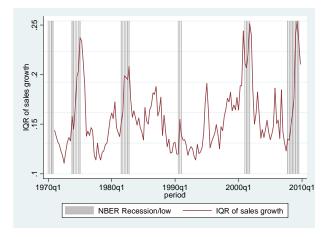
#### Recent recession

- ▶ Output and labor drop, accounted for
  - ▶ Mainly by a worsening of labor wedge
  - Less by a fall in TFP

#### Popular story

- ▶ Increase in "uncertainty" at firm level
- ▶ Interacts with financial frictions
- $\Rightarrow$  Firms shrink level of employment

#### Interquartile range sales growth 1970-2010



- Dispersion has increased in many recessions
- ▶ Not arguing that recessions driven solely by uncertainty shocks

#### Key Elements in Model

- ▶ Firms produce before knowing current idiosyncratic demand shock
  - ▶ In high states 'too small' and in low states 'too big'
- ▶ Firms have limited ability to insure idiosyncratic shock
  - ▶ If scale too big, can't pay wage bill and might default
- ▶ Costly default
  - ▶ Liquidated, so lose future profits that are covering entry cost
- $\Longrightarrow$  Labor wedge
  - ▶ Risk of default create a wedge between MPL and wage

Volatility shock generates labor-wedge driven recession

▶ Increase in volatility

- ▶ Increases risk of default for a given scale
- ▶ Induces firms to choose smaller scale
- ▶ So increases wedge between MPL and wage

## Simple Example

## Simple Example

#### ▶ Period 1:

- $\blacktriangleright$  Firms hire labor and produce before the productivity shock z
- ▶ Firms are liquidated if dividend is negative
- ▶ Period 2:
  - $\blacktriangleright$  Firms get future value V only if not liquidated

#### Complete financial markets

 $\blacktriangleright$  Firms choose  $\ell$  to maximize the expected value

$$\max_{\ell} \int_0^\infty \left[ z\ell^\theta - w\ell + V \right] f(z) dz$$

▶ Optimal scale chosen to maximize short term profits

$$\underbrace{\theta \ell^{\theta - 1} Ez}_{\text{value } MPL} = w$$

▶ Use state-contingent debt to pay dividends and avoid liquidation

#### Incomplete financial markets

▶ Firms are liquidated when demand shocks are low  $(z < \hat{z})$ 

For each  $\ell$ ,  $\hat{z}$  is lowest z s.t.  $z\ell^{\theta} \ge w\ell$ 

Firms choose  $(\ell, \hat{z})$  to maximize the expected value

$$\max_{\ell,\hat{z}} \int_{\hat{z}}^{\infty} \left[ z\ell^{\theta} - w\ell \right] f(z)dz + \int_{\hat{z}}^{\infty} Vf(z)dz$$
s.t
$$z\ell^{\theta} - w\ell = 0$$

 Optimal scale chosen to maximize short term profits and future value

$$\underbrace{\theta \ell^{\theta-1} E[z|z \ge \hat{z}]}_{MPL} = w + \underbrace{V \frac{f(\hat{z})}{1 - F(\hat{z})} \frac{d\hat{z}}{d\ell}}_{\text{Wedge}}$$

#### Incomplete financial markets

$$\underbrace{\theta \ell^{\theta-1} E[z|z \ge \hat{z}]}_{MPL} = w + \underbrace{V \frac{f(\hat{z})}{1 - F(\hat{z})} \frac{d\hat{z}}{d\ell}}_{\text{Wedge}}$$

- ▶ Optimal labor may higher than the efficient case since  $E(z|z \ge \hat{z}) \ge Ez$
- $\blacktriangleright$  When V is large enough, incomplete market has lower scale than the complete market model
- Increased volatility reduces labor and output and worsens labor wedge

# Model

#### Our model

Dynamic general equilibrium model with

- ► Households (standard)
  - Provide labor
  - Sell uncontingent debt to firms
  - ► Own firms
- ▶ Final goods firms
  - ▶ Aggregate intermediate goods with CES aggregator
- ► Firms

#### Financial frictions

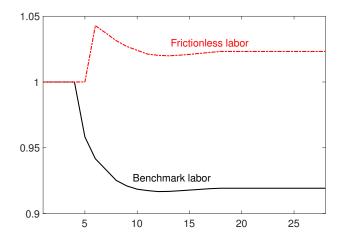
- ▶ Limited liability, dividend has to be nonnegative
- ▶ Firm can only borrow state uncontingent bond, but can default
  - ▶ Defaulting firms receive zero value and exit
  - ▶ If output is less than labor costs, the government pays the gap
- ▶ Agency friction as in Jensen (1986)
  - Manager may secretly borrow the unused credit and use for his own benefit
  - Shareholders understand this incentive of diverting and prefer overborrowing

# Experiments and Results

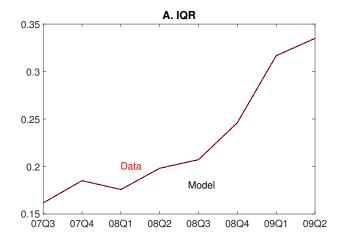
## Quantifying volatility shocks

- ▶ Use cross-section firm dispersion to parameterize volatility shocks
- ► As in Bloom (2009), we restrict the sample for firms to those with at least 100 quarters of observations since 1970
- ▶ Firm dispersion:
  - $\blacktriangleright$  Interquartile range of sales growth (differences between 75% and 25%)

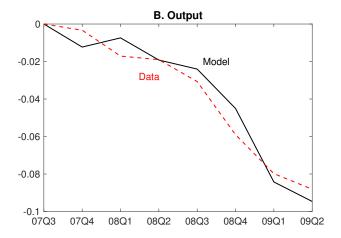
#### Firm Labor Impulse Responses



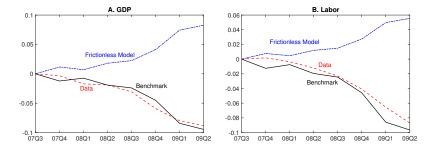
#### Great Recession Event



#### Great Recession Event



#### Great Recession Event, Frictionless Financial Markets



#### Conclusion

- Framework that combines volatility shocks with financial markets imperfections
- ▶ Generates movements in output, labor, and the labor wedge linked to financial frictions